Multilevel Monte Carlo for PDE solutions based on Feynman-Kac theorem

Mike Giles, Francisco Bernal

Mathematical Institute, University of Oxford, UK

Instituto Superior Tecnico, Portugal

ICIAM 2015

August 10, 2015

Outline

- Feynman-Kac formula
- prior work Gobet & Menozzi
- multilevel Monte Carlo
- prior work Higham et al
- new idea approximating a conditional expectation
- outline numerical analysis

Feynman-Kac formula

Suppose that u(x, t) satisfies the parabolic PDE

$$\frac{\partial u}{\partial t} + \sum_{j} a_{j} \frac{\partial u}{\partial x_{j}} + \frac{1}{2} \sum_{j,k,l} b_{jl} b_{kl} \frac{\partial^{2} u}{\partial x_{j} \partial x_{k}} - V u + f = 0$$

in bounded domain D, subject to u(x,t) = g(x,t) on the boundary ∂D .

It will be assumed that f(x,t), V(x,t), a(x,t), b(x,t) are all Lipschitz continuous, and g(x,t) is continuously twice-differentiable.

Feynman-Kac formula

Feynman and Kac proved that u(x,t) can also be expressed as

$$u(x,t) = \mathbb{E}\left[\int_t^{\tau} E(t,s) f(X_s,s) ds + E(t,\tau) g(X_{\tau},\tau) \mid X_t = x\right]$$

where X_t satisfies the SDE

$$\mathrm{d}X_t = a(X_t,t)\,\mathrm{d}t + b(X_t,t)\,\mathrm{d}W_t,$$

with W_t being a Brownian motion with independent components, τ is the first time at which X_t leaves D, and

$$E(t_0,t_1)=\exp\left(-\int_{t_0}^{t_1}V(X_t,t)\,\mathrm{d}t\right).$$

Note: in the special case in which f(x,t)=0, g(x,t)=t, V(x,t)=0 $u_{exit}(x,t)$ is the expected exit time.

<ロ > < 回 > < 回 > < 目 > < 目 > < 目 > の < や

Numerical approximation

An Euler-Maruyama discretisation with uniform timestep h gives

$$\widehat{X}_{t_{n+1}} = \widehat{X}_{t_n} + a(\widehat{X}_{t_n}, t_n) h + b(\widehat{X}_{t_n}, t_n) \Delta W_n,$$

with initial data $\widehat{X}_0 = x$ at time t.

If $\widehat{X}(t)$ is the piecewise-constant interpolant, we then have

$$\widehat{u}(x,t) = \mathbb{E}\left[\int_t^{\widehat{\tau}} \widehat{E}(t,s) f(\widehat{X}(s),s) ds + \widehat{E}(t,\widehat{\tau}) g(\widehat{X}(\tau),\widehat{\tau})\right].$$

with $\widehat{ au}$ being the exit time, and

$$\widehat{E}(t_0,t_1) = \exp\left(-\int_{t_0}^{t_1} V(\widehat{X}_t,t) dt\right).$$

◆ロト ◆団ト ◆草ト ◆草ト ■ りゅぐ

Prior work - Gobet & Menozzi

The Euler-Maruyama method has strong accuracy

$$\left(\mathbb{E}\left[\sup_{[0,\min(\tau,\widehat{\tau})]}\|X_t - \widehat{X}(t)\|^2\right]\right)^{1/2} = O(h^{1/2}|\log h|^{1/2}),$$

and Gobet & Menozzi (2007) proved that it has weak error

$$u(x,t)-\widehat{u}(x,t)=O(h^{1/2}).$$

For standard Monte Carlo method, ε RMS accuracy needs $O(\varepsilon^{-2})$ paths, each with $h=O(\varepsilon^2)$, so total cost is $O(\varepsilon^{-4})$

Gobet & Menozzi (2010) reduced this to $O(\varepsilon^{-3})$ by shifting the boundary by $O(h^{1/2})$ to improve the weak error to O(h).

Multilevel Monte Carlo

MLMC is based on the telescoping sum

$$\mathbb{E}[\widehat{P}_L] = \mathbb{E}[\widehat{P}_0] + \sum_{\ell=1}^L \mathbb{E}[\widehat{P}_\ell - \widehat{P}_{\ell-1}]$$

where \widehat{P}_ℓ represents an approximation using timestep $h_\ell=2^{-\ell}\,h_0$, with weak convergence

$$\mathbb{E}[\widehat{P}_{\ell} - P] = O(2^{-\alpha \, \ell})$$

If \widehat{Y}_{ℓ} is an unbiased estimator for $\mathbb{E}[\widehat{P}_{\ell}-\widehat{P}_{\ell-1}]$, based on N_{ℓ} samples, with variance

$$\mathbb{V}[\widehat{Y}_{\ell}] = O(N_{\ell}^{-1} \, 2^{-\beta \, \ell})$$

and expected cost

$$\mathbb{E}[C_{\ell}] = O(N_{\ell} 2^{\gamma \ell}), \ldots$$

< □ > < □ > < 亘 > < 亘 > □ ■ ● へ ○ ○

Multilevel Monte Carlo

... then the finest level L and the number of samples N_{ℓ} on each level can be chosen to achieve an RMS error of ε at an expected cost

$$C = \left\{ \begin{array}{ll} O\left(\varepsilon^{-2}\right), & \beta > \gamma, \\ \\ O\left(\varepsilon^{-2}(\log \varepsilon)^2\right), & \beta = \gamma, \\ \\ O\left(\varepsilon^{-2-(\gamma-\beta)/\alpha}\right), & 0 < \beta < \gamma. \end{array} \right.$$

Prior work – Higham

Higham *et al* (2013) developed a MLMC treatment of the exit time problem:

- Euler-Maruyama discretisation
- $O(h_{\ell}^{1/2})$ weak convergence $\implies \alpha = 1/2$
- $\mathbb{V}[\widehat{P}_{\ell} \widehat{P}_{\ell-1}] = O(h_{\ell}^{1/2} |\log h_{\ell}|^{1/2}) \Longrightarrow \beta \approx 1/2$
- ullet $O(h_\ell^{-1})$ cost per path $\Longrightarrow \ \gamma=1$

Hence, overall cost is $O(\varepsilon^{-3}|\log \varepsilon|^{1/2})$.

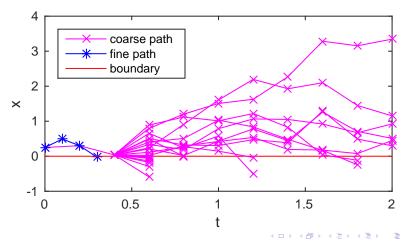
Gobet & Menozzi's boundary treatment would improve this to $O(\varepsilon^{-2.5}|\log \varepsilon|^{1/2})$.

G & Primozic (2011) developed $O(\varepsilon^{-2})$ treatment using Milstein discretisation for SDEs with special commutativity property.

MLMC challenge

When coarse or fine path exits the domain, the other is within $O(h^{1/2})$.

However, there is a $O(h^{1/2})$ probability that it will not exit the domain until much later $\Longrightarrow V_\ell \approx O(h^{1/2})$.



MLMC challenge

How can we do better?

Similar to previous work on digital options (G, Burgos), and also used by Dickmann & Schweizer for stopping times, split second path into multiple copies and average their outputs to approximate the conditional expectation.

Approximately $O(h^{1/2})$ expected time to exit for second path, so can afford to use approximately $O(h^{-1/2})$ copies of second path.

This gives an approximation to the conditional expectation resulting in $\widehat{P}_{\ell}-\widehat{P}_{\ell-1}\approx O(h^{1/2})$, so $V_{\ell}\approx O(h)$.

This gives $\alpha=1/2,\ \beta\approx 1,\ \gamma\approx 1$ and the complexity is $O(\varepsilon^{-2}\,|\log\varepsilon|^3)$.

Assumption 1: There is a Lipschitz constant L_f such that

$$|f(x,t)-f(y,s)| \le L_f(||x-y||_2+|t-s|), \quad \forall (x,t), (y,s) \in D,$$

and there are similar Lipschitz constants L_g , L_V , L_a , L_b , L_u , L_{exit} for g,V,a,b,u,u_{exit} . In addition, $g\in C^{2,1}(D)$, with a bounded Hessian $H_g\equiv \partial^2 g/\partial x^2$.

Comment: assumption about L_u , L_{exit} may require the boundary ∂D to be smooth, or at least not have re-entrant corners.

Assumption 2: There is a unit computational cost for each timestep, and in determining whether or not $\widehat{X}_{t_{n+1}} \in D$.

Assumption 3: There exist constants C_u and C_{exit} s.t. for all $(x,t) \in D$

$$|u(x,t) - \widehat{u}(x,t)| \leq C_u h^{1/2}$$

$$|u_{exit}(x,t) - \widehat{u}_{exit}(x,t)| \leq C_{exit} h^{1/2}$$

Defining the output functional

$$P_t = \int_t^{\tau} E(t,s) f(X_s,s) ds + E(t,\tau) g(X_{\tau},\tau)$$

we get

Lemma

Given Assumption 1, there exists C such that for any $(x,t) \in D$

$$\mathbb{V}[P_t \mid X_t = x] \leq C \mathbb{E}[\tau - t \mid X_t = x].$$

$$d\left(E(t,s)g(X_s,s)\right) = \\ E(t,s)\left(\left(-Vg + \dot{g} + (\nabla g)^T a + \frac{1}{2}\operatorname{trace}(b^T H_g b)\right) ds + (\nabla g)^T b dW_s\right)$$

with a, b, g, $\dot{g} \equiv \partial g/\partial t$, ∇g , H_g , all evaluated at (X_s, s) .

Hence, $P_t - g(x, t) = p^{(1)} + p^{(2)}$, where

$$p^{(1)} = \int_{t}^{\tau} E(t,s) \left(f - V g + \dot{g} + (\nabla g)^{T} a + \frac{1}{2} \operatorname{trace}(b^{T} H_{g} b) \right) ds,$$

$$p^{(2)} = \int_{t}^{\tau} E(t,s) (\nabla g)^{T} b dW_{s}.$$

Considering the second term, since $E(t,s) \leq \exp(T\|V\|_{\infty})$, we have

$$\mathbb{E}[(p^{(2)})^{2}] = \mathbb{E}\left[\int_{t}^{\tau} (E(t,s))^{2} \|(\nabla g)^{T} b\|_{2}^{2} ds\right]$$

$$\leq \exp(2T\|V\|_{\infty}) \|\nabla g\|_{2,\infty}^{2} \|b\|_{2,\infty}^{2} \mathbb{E}[\tau - t | X_{t} = x],$$

where $\|b\|_{2,\infty}$, $\|\nabla g\|_{2,\infty}$ are the maximum values of $\|b\|_2$, $\|\nabla g\|_2$ over D.

The first term is handled similarly to complete the proof.

◆ロ> ◆部> ◆注> ◆注> 注 のQで

The following is a standard result:

Lemma

If W and Z are independent random variables, then

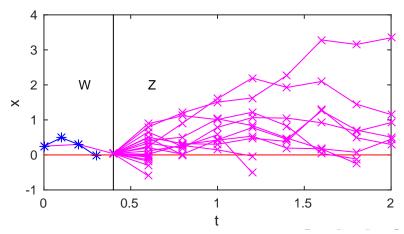
$$Y = M^{-1} \sum_{m=1}^{M} f(W, Z^{(m)})$$

with independent samples W and $Z^{(m)}$ is an unbiased estimator for $\mathbb{E}\left[f(W,Z)\right]$ and its variance is

$$\mathbb{V}[Y] = \mathbb{V}\left[\mathbb{E}[f(W,Z)\,|\,W]\right] + M^{-1}\,\mathbb{E}\left[\mathbb{V}[f(W,Z)\,|\,W]\right].$$

Let $\underline{\tau}$ be the exit time of the first of a pair of coarse/fine paths, and $\overline{\tau}$ be $\underline{\tau}$ rounded up to the end of a coarse timestep.

In our application W represents the Brownian path up to $\overline{\tau}$, and Z is the Brownian path therafter.



Lemma

Given Assumptions 1 and 3, we have

$$\mathbb{E}[\sup_{[0,\underline{\tau}]} \|\widehat{X}_{\ell,t} - \widehat{X}_{\ell-1,t}\|^{2}]^{1/2} = O(h_{\ell-1}^{1/2} |\log h_{\ell-1}|^{1/2})$$

$$\mathbb{E}[\|\widehat{X}_{\ell,\underline{\tau}} - \widehat{X}_{\ell-1,\overline{\tau}}\|^{2}]^{1/2} = O(h_{\ell-1}^{1/2} |\log h_{\ell-1}|^{1/2})$$

$$\Longrightarrow \mathbb{V}\left[\mathbb{E}[\widehat{P}_{\ell} - \widehat{P}_{\ell-1} | W]\right] = O(h_{\ell-1} |\log h_{\ell-1}|)$$

The key to the proof is that if $0 \le t \le \tau$ then

$$P_{0} = \int_{0}^{t} E(0,s) f(X_{s},s) ds + E(0,t) \left\{ \int_{t}^{\tau} E(t,s) f(X_{s},s) ds + E(t,\tau) g(X_{\tau},\tau) \right\}$$

$$\Rightarrow \mathbb{E}[P_{0} | \mathcal{F}_{t}] - \int_{0}^{t} E(0,s) f(X_{s},s) ds = E(0,t) \mathbb{E}[P_{t} | \mathcal{F}_{t}] = E(0,t) u(X_{t},t)$$

Something similar for the discrete approximation yields

$$\mathbb{E}[\widehat{P}_{\ell} - \widehat{P}_{\ell-1} \mid W] = O(h_{\ell-1}^{1/2} |\log h_{\ell-1}|^{1/2})$$

Lemma

Given Assumptions 1 and 3,

$$\mathbb{E}\left[\mathbb{V}[\widehat{P}_{\ell} - \widehat{P}_{\ell-1} \,|\, W]\right] = O(h_{\ell-1}^{1/2} |\log h_{\ell-1}|^{1/2})$$

The key here is that, similar to the SDE analysis, there exists C such that

$$\mathbb{V}[\widehat{P}_{\ell} - \widehat{P}_{\ell-1} \mid W] \leq C \mathbb{E}[|\widehat{\tau}_{\ell} - \widehat{\tau}_{\ell-1}| \mid W])
= O(h_{\ell-1}^{1/2} |\log h_{\ell-1}|^{1/2})$$

Corollary

Under the given assumptions, an RMS error of ε can be achieved with an $O(\varepsilon^{-2}|\log \varepsilon|^3)$ expected computational cost.

The proof is slightly non-standard because of log terms.

- $h_{\ell} = 4^{-\ell} h_0$
 - $M_\ell = \lceil 2^\ell/\ell^{1/2} \rceil$ paths in the splitting estimator
 - expected cost is $O(h_{\ell}^{-1})$
 - variance $V_\ell = O(h_\ell | \log h_\ell|) = O(h_\ell \ell)$.

This eventually gives the desired cost bound.

Conclusions

- conditional expectation / splitting is a useful technique in MLMC estimation
- in Feynmac-Kac application it improves the MLMC variance from approximately $O(h^{1/2})$ to approximately O(h), reducing the complexity to $O(\varepsilon^{-2}|\log \varepsilon|^3)$
- numerical analysis is now complete but relies on key assumption of uniform $O(h^{1/2})$ weak convergence an open problem

Webpages:

```
people.maths.ox.ac.uk/gilesm/mlmc.html
people.maths.ox.ac.uk/gilesm/mlmc_community.html
people.maths.ox.ac.uk/gilesm/acta/
```