"Vibrato" Monte Carlo evaluation of Greeks

(Smoking Adjoints: part 3)

Mike Giles

mike.giles@maths.ox.ac.uk

Oxford University Mathematical Institute
Oxford-Man Institute of Quantitative Finance

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"Smoking Adjoints"

Paper with Paul Glasserman in *Risk* in 2006 showed how adjoints can be used in computing pathwise sensitivities – gives lots of first order sensitivities for negligible cost

This attracted a lot of interest, and questions:

- what is involved in practice in creating an adjoint code, and can it be simplified? (see HERCMA paper, available from website)
- do we really have to differentiate the payoff?
- what about discontinuous payoffs?
- what about American options? (not addressed yet!)

Outline

- different approaches to computing Greeks
 - finite differences
 - likelihood ratio method
 - pathwise sensitivity
- use of conditional expectation for a digital option
- "vibrato" extension for scalar SDE
- generalisation to multidimensional SDEs

Generic Problem

Stochastic differential equation with general drift and volatility terms:

$$dS_t = a(S_t, t) dt + b(S_t, t) dW_t$$

For a simple European option we want to compute the expected discounted payoff value dependent on the terminal state:

$$V = \mathbb{E}[f(S_T)]$$

Note: the drift and volatility functions are almost always differentiable, but the payoff f(S) is often not.

Generic Problem

Euler discretisation with timestep *h*:

$$\widehat{S}_{n+1} = \widehat{S}_n + a(\widehat{S}_n, t_n) h + b(\widehat{S}_n, t_n) \Delta W_n$$

Simplest Monte Carlo estimator for expected payoff is an average of M independent path simulations:

$$M^{-1} \sum_{i=1}^{M} f(\widehat{S}_N^{(i)})$$

Greeks: for hedging and risk management we also want to estimate derivatives of expected payoff ${\cal V}$

Simple Problem

For Geometric Brownian motion

$$dS_t = r S_t dt + \sigma S_t dW_t$$

the SDE can be solved analytically to give

$$S_T = S_0 \exp\left(\left(r - \frac{1}{2}\sigma^2\right)T + \sigma W_T\right)$$

In this case, we can directly sample W_T to get

$$V \equiv \mathbb{E}\left[f(S_T)\right] \approx M^{-1} \sum_{i=1}^{M} f(S_T^{(i)})$$

 will use this to illustrate approaches to calculating sensitivities

Finite Differences

Simplest approach is to use a finite difference approximation,

$$\frac{\partial V}{\partial \theta} \approx \frac{V(\theta + \Delta \theta) - V(\theta - \Delta \theta)}{2 \Delta \theta}$$

$$\frac{\partial^2 V}{\partial \theta^2} \approx \frac{V(\theta + \Delta \theta) - 2V(\theta) + V(\theta - \Delta \theta)}{(\Delta \theta)^2}$$

– very simple, but expensive and inaccurate if $\Delta\theta$ is too big, or too small in the case of discontinuous payoffs

Likelihood Ratio Method

For simple cases where we know the terminal probability distribution

$$V \equiv \mathbb{E}\left[f(S_T)\right] = \int f(S) \ p_S(\theta; S) \ dS$$

we can differentiate this to get

$$\frac{\partial V}{\partial \theta} = \int f \frac{\partial p_S}{\partial \theta} dS = \int f \frac{\partial (\log p_S)}{\partial \theta} p_S dS = \mathbb{E} \left[f \frac{\partial (\log p_S)}{\partial \theta} \right]$$

This is the Likelihood Ratio Method (Broadie & Glasserman, 1996) – its great strength is that it can handle discontinuous payoffs

Likelihood Ratio Method

The LRM weakness is in its generalisation to full path simulations for which we get the multi-dimensional integral

$$\widehat{V} = \mathbb{E}[f(\widehat{S})] = \int f(\widehat{S}) p(\widehat{S}) d\widehat{S},$$

where

$$d\widehat{S} \equiv d\widehat{S}_1 \ d\widehat{S}_2 \ d\widehat{S}_3 \ \dots \ d\widehat{S}_N$$

and the joint probability density function $p(\widehat{S})$ is the product of the p.d.f.s for each timestep

$$p(\widehat{S}) = \prod_{n} p_n(\widehat{S}_{n+1}|\widehat{S}_n)$$
$$\log p(\widehat{S}) = \sum_{n} \log p_n(\widehat{S}_{n+1}|\widehat{S}_n)$$

Likelihood Ratio Method

When computing Vega from an Euler discretisation of Geometric Brownian motion this leads to

$$\frac{\partial \widehat{V}}{\partial \sigma} = \mathbb{E}\left[\left(\sum_{n} \frac{Z_{n}^{2} - 1}{\sigma}\right) f(\widehat{S}_{N})\right]$$

where Z_n is the unit Normal used in the n^{th} timestep

$$\widehat{S}_{n+1} = \widehat{S}_n(1+rh) + \sigma \,\widehat{S}_n \,\sqrt{h} \,Z_n$$

Since $\mathbb{V}[Z_n^2-1]=2$ it follows that the variance of the estimator is $O(h^{-1})$

This blow-up as $h \rightarrow 0$ is the weakness of the LRM.

Pathwise sensitivities

Alternatively, for simple Geometric Brownian Motion

$$V \equiv \mathbb{E}\left[f(S_T)\right] = \int f(S_T(\theta; W)) \ p_W(W) \ dW$$

and differentiating this gives

$$\frac{\partial V}{\partial \theta} = \int \frac{\partial f}{\partial S} \frac{\partial S_T}{\partial \theta} p_W dW = \mathbb{E} \left[\frac{\partial f}{\partial S} \frac{\partial S_T}{\partial \theta} \right]$$

with $\partial S_T/\partial \theta$ being evaluated at fixed W.

This is the pathwise sensitivity approach – it can't handle discontinuous payoffs, but generalises well to full path simulations

Pathwise sensitivities

The generalisation involves differentiating the Euler path discretisation,

$$\widehat{S}_{n+1} = \widehat{S}_n + a(\widehat{S}_n, t_n) h + b(\widehat{S}_n, t_n) \Delta W_n$$

holding fixed the Brownian increments, to get

$$\frac{\partial \widehat{S}_{n+1}}{\partial \theta} = \left(1 + \frac{\partial a}{\partial S}h + \frac{\partial b}{\partial S}\Delta W_n\right)\frac{\partial \widehat{S}_n}{\partial \theta} + \frac{\partial a}{\partial \theta}h + \frac{\partial b}{\partial \theta}\Delta W_n$$

leading to

$$\frac{\partial \widehat{V}}{\partial \theta} = \mathbb{E} \left[\frac{\partial f}{\partial S} (\widehat{S}_N) \frac{\partial \widehat{S}_N}{\partial \theta} \right].$$

Pathwise sensitivities

In the case of Vega for an Euler discretisation of GBM

$$\widehat{S}_{n+1} = \widehat{S}_n + r\,\widehat{S}_n\,h + \sigma\,\widehat{S}_n\,\Delta W_n$$

we get

$$\frac{\partial \widehat{S}_{n+1}}{\partial \sigma} = \left(1 + rh + \sigma \Delta W_n\right) \frac{\partial \widehat{S}_n}{\partial \sigma} + \widehat{S}_n \Delta W_n$$

and the variance

$$\mathbb{V}\left[\frac{\partial f}{\partial S}(\widehat{S}_N) \frac{\partial \widehat{S}_N}{\partial \sigma}\right]$$

is O(1) if f(S) is Lipschitz.

What is best if payoff is discontinuous?

- LRM
 - estimator variance $O(h^{-1})$
- Malliavin calculus
 - estimator variance O(1)
 - recent paper by Glasserman & Chen shows it can be viewed as a pathwise/LRM hybrid
 - might be good choice when few Greeks needed
- new "vibrato" Monte Carlo idea
 - also a pathwise/LRM hybrid
 - estimator variance $O(h^{-1/2})$
 - efficient adjoint implementation

- new idea is based on use of conditional expectation for a simple digital option in Paul Glasserman's book
- output of each SDE path calculation becomes a narrow (multivariate) Normal distribution
- combine pathwise sensitivity for the differentiable SDE, with LRM for the discontinuous payoff
- avoiding the differentiation of the payoff also simplifies the implementation in real-world setting

Final timestep of Euler path discretisation is

$$\widehat{S}_N = \widehat{S}_{N-1} + a(\widehat{S}_{N-1}, t_{N-1}) h + b(\widehat{S}_{N-1}, t_{N-1}) \Delta W_{N-1}$$

Instead of using random number generator to get a value for ΔW_{N-1} , consider the whole distribution of possible values, so \widehat{S}_N has a Normal distribution with mean

$$\mu_W = \widehat{S}_{N-1} + a(\widehat{S}_{N-1}, t_{N-1}) h$$

and standard deviation

$$\sigma_W = b(\widehat{S}_{N-1}, t_{N-1}) \sqrt{h}$$

where $W \equiv (\Delta W_0, \Delta W_1, \dots \Delta W_{N-2})$.

For a particular path given by a particular vector W, the expected payoff is

$$\mathbb{E}_{Z}[f(\mu_{W} + \sigma_{W} Z)]$$

where Z is a unit Normal random variable.

Averaging over all W then gives the same overall expectation as before.

Note also that, for given W, \widehat{S}_N has a Normal distribution

$$p_S(\widehat{S}) = \frac{1}{\sqrt{2\pi} \,\sigma_W} \,\exp\left(-\frac{(\widehat{S} - \mu_W)^2}{2 \,\sigma_W^2}\right)$$

In the case of a simple digital call with strike K, the analytic solution is

$$\mathbb{E}_{Z}[f(\mu_{W} + \sigma_{W} Z)] = \exp(-rT) \Phi\left(\frac{\mu_{W} - K}{\sigma_{W}}\right).$$

- ullet for each W, the payoff is now smooth, differentiable
- derivative is $O(h^{-1/2})$ near strike, near zero elsewhere \Longrightarrow variance is $O(h^{-1/2})$
- analytic evaluation of conditional expectation not possible in general for multivariate cases
 use Monte Carlo estimation!

Main novelty comes in calculating the sensitivity.

For a particular W, we have a Normal probability distribution for \widehat{S}_N and can apply the Likelihood Ratio method to get

$$\frac{\partial}{\partial \theta} \mathbb{E}_Z \left[f(\widehat{S}_N) \right] = \mathbb{E}_Z \left[f(\widehat{S}_N) \frac{\partial (\log p_S)}{\partial \theta} \right],$$

where

$$\frac{\partial(\log p_S)}{\partial \theta} = \frac{\partial(\log p_S)}{\partial \mu_W} \frac{\partial \mu_W}{\partial \theta} + \frac{\partial(\log p_S)}{\partial \sigma_W} \frac{\partial \sigma_W}{\partial \theta} \\
= \frac{Z}{\sigma_W} \frac{\partial \mu_W}{\partial \theta} + \frac{Z^2 - 1}{\sigma_W} \frac{\partial \sigma_W}{\partial \theta}.$$

Averaging over all W then gives the expected sensitivity.

To improve the variance, we note that

$$\mathbb{E}_{Z} \left[f(\mu_{W} + \sigma_{W} Z) Z \right] = \mathbb{E}_{Z} \left[-f(\mu_{W} - \sigma_{W} Z) Z \right]$$

$$= \frac{1}{2} \mathbb{E}_{Z} \left[\left(f(\mu_{W} + \sigma_{W} Z) - f(\mu_{W} - \sigma_{W} Z) \right) Z \right]$$

and similarly

$$\begin{split} &\mathbb{E}_{Z}\left[f(\mu_{W}+\sigma_{W}Z)\;(Z^{2}-1)\;\right]\\ &=\; \frac{1}{2}\;\mathbb{E}_{Z}\left[\left(f(\mu_{W}+\sigma_{W}Z)-2f(\mu_{W})+f(\mu_{W}-\sigma_{W}Z)\right)\;(Z^{2}-1)\;\right] \end{split}$$

This gives an estimator with O(1) variance when f(S) is Lipschitz, and $O(h^{-1/2})$ variance when it is discontinuous.

Test case: Geometric Brownian motion

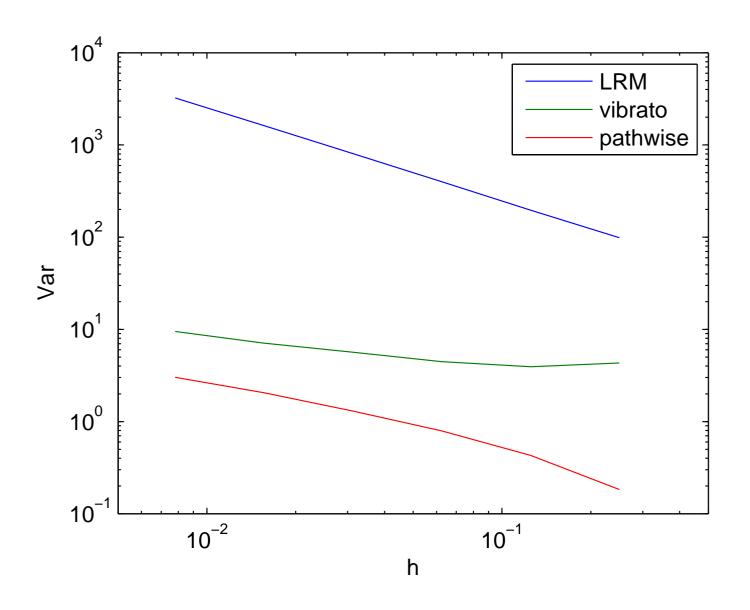
$$dS_t = r S_t dt + \sigma S_t dW_t$$

with simple digital call option.

Parameters: $r = 0.05, \ \sigma = 0.2, \ T = 1, \ S_0 = 100, \ K = 100$

Numerical results compare:

- LRM
- ullet vibrato with one Z per W
- pathwise with conditional expectation



These results used just one Z per path. If M_Z are used, the variance is

$$\mathbb{V}_W\Big[\mathbb{E}_Z[g(W,Z)]\Big] + M_Z^{-1}\mathbb{E}_W\Big[\mathbb{V}_Z[g(W,Z)]\Big]$$

where g(W,Z) is the estimator.

The limit $M_z \to \infty$ gives the variance for the estimator based on the analytic conditional expectation.

The optimal M_Z can be determined if one knows/estimates $\mathbb{V}_W\big[\mathbb{E}_Z[g(W,Z)]\big]$ and $\mathbb{E}_W\big[\mathbb{V}_Z[g(W,Z)]\big]$, and the relative cost of the path simulation and the payoff evaluation.

Multivariate extension

In general we have

$$\widehat{S}(W,Z) = \mu_W + C_W Z$$

where $\Sigma_W = C_W C_W^T$ is the covariance matrix, and Z is a vector of uncorrelated Normals. The joint p.d.f. is

$$\log p_S = -\frac{1}{2} \log |\Sigma_W| - \frac{1}{2} (\widehat{S} - \mu_W)^T \Sigma_W^{-1} (\widehat{S} - \mu_W) - \frac{1}{2} d \log(2\pi)$$

and so

$$\frac{\partial \log p_S}{\partial \mu_W} = C_W^{-T} Z,$$

$$\frac{\partial \log p_S}{\partial \Sigma_W} = \frac{1}{2} \, C_W^{-T} \left(Z Z^T - I \right) C_W^{-1}$$

Multivariate extension

This leads to

$$\frac{\partial}{\partial \theta} \mathbb{E}_Z \left[f(\widehat{S}) \right] = \mathbb{E}_Z \left[f(\widehat{S}) \frac{\partial (\log p_S)}{\partial \theta} \right]$$

where

$$\frac{\partial (\log p_S)}{\partial \theta} = \left(\frac{\partial \log p_S}{\partial \mu_W}\right)^T \frac{\partial \mu_W}{\partial \theta} + \operatorname{tr}\left(\frac{\partial \log p_S}{\partial \Sigma_W} \frac{\partial \Sigma_W}{\partial \theta}\right)$$

and $\frac{\partial \mu_W}{\partial \theta}$, $\frac{\partial \Sigma_W}{\partial \theta}$ come from pathwise sensitivity analysis.

A more efficient estimator can be obtained by similar reasoning to the scalar case.

Test case: Geometric Brownian motion

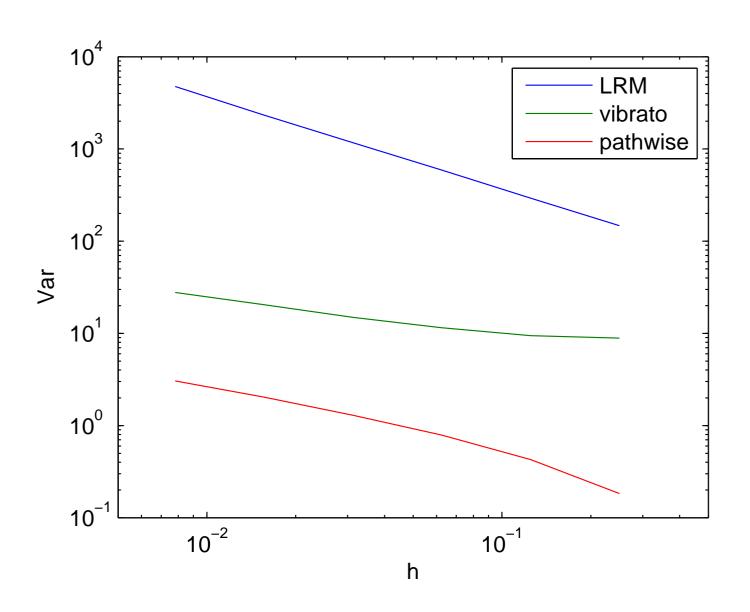
$$dS_t^{(1)} = r S_t^{(1)} dt + \sigma^{(1)} S_t^{(1)} dW_t^{(1)}$$

$$dS_t^{(1)} = r S_t^{(2)} dt + \sigma^{(2)} S_t^{(2)} dW_t^{(2)}$$

with a simple digital call option based solely on $S_T^{(1)}$.

Parameters:
$$r=0.05,\ \sigma^{(1)}=0.2,\ \sigma^{(2)}=0.3,\ T=1,\ S_0^{(1)}=S_0^{(2)}=100,\ K=100,\ \rho=0.5$$

Numerical results again compare LRM, vibrato with one Z per W, and pathwise with conditional expectation.



Multivariate extension

Can also treat payoffs dependent on $S(\tau)$ at intermediate times, by taking

$$t_n < \tau < t_{n+1}$$

and using simple Brownian motion interpolation between \widehat{S}_n and \widehat{S}_{n+1} to get a Normal distribution for $\widehat{S}(\tau)$, with

mean:

$$\widehat{S}_n + \frac{\tau - t_n}{t_{n+1} - t_n} \left(\widehat{S}_{n+1} - \widehat{S}_n \right)$$

variance:

$$\frac{(\tau - t_n)(t_{n+1} - \tau)}{t_{n+1} - t_n} b^2(\widehat{S}_n, t_n)$$

Conclusions

"Vibrato" idea for computing Greeks offers

- O(1) variance for Lipschitz payoffs, and easy implementation no derivatives required
- $O(h^{-1/2})$ variance for discontinuous payoffs
- adjoint implementation for multiple Greeks

Future work:

similar idea for digital options in multilevel Monte Carlo path simulation – introduces Radon-Nikodym derivative from change in measure

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Further information

- www.maths.ox.ac.uk/~gilesm/
- Email: mike.giles@maths.ox.ac.uk