

Predictability of Large Future Changes in a Competitive Evolving Population

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The dynamical evolution of many economic, sociological, biological, and physical systems tends to be dominated by a relatively small number of unexpected, large changes (“extreme events”). We study the large, internal changes produced in a generic multiagent population competing for a limited resource, and find that the level of predictability *increases* prior to a large change. These large changes hence arise as a predictable consequence of information encoded in the system’s global state.

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Populations comprising many “agents” (e.g., people, species, data packets, cells) who compete for a limited resource are believed to underlie the complex dynamics observed in areas as diverse as economics [1–4], sociology [5], internet traffic [6], ecology [7], and biology [8,9]. The reliable prediction of large future changes (“extreme events”) in such complex systems would be of enormous practical importance, but is widely considered to be impossible [10].

In this paper, we examine the predictability of large future changes produced within an evolving population of agents who compete for a limited resource. We find that the level of predictability in the system *increases* prior to a large change, implying that such a large change arises as a predictable consequence of information encoded in the system’s global state, as opposed to being triggered by some isolated random event.

We consider a generic multiagent system comprising a population of N_{tot} agents of which no more than $L < N_{\text{tot}}$ agents can win at each time step; an everyday example would be a popular bar with a limited seating capacity L [5]. For the purpose of this paper, we consider a specific case of such a limited-resource problem with N_{tot} odd and $L = (N_{\text{tot}} - 1)/2$ [11], hence there are more losers than winners, noting that similar dynamics can also occur for more general L [12]. Each agent is therefore seeking to be in the minority group: for example, a buyer in a financial market may obtain a better price if more people are selling than buying; a driver may have a quicker journey if she chooses the route with less traffic. At each time step, an agent decides whether to enter a game where the choices are option 0 (e.g., buy, choose route A) and option 1 (e.g., sell, choose route B). Each agent holds a finite number of strategies and only a subset $N = N_0 + N_1 \leq N_{\text{tot}}$ of the population, who are sufficiently confident of winning, actually play: N_0 agents choose 0 while N_1 choose 1. If $N_0 - N_1 > 0$, the winning decision (outcome) is “1” and vice versa. If $N_0 = N_1$ the tie is decided by a coin toss. Hence N and the “excess demand” $N_{0-1} = N_0 - N_1$ both fluctuate with time. In contrast to the basic minority game

(MG) [11], this variable- N model has the realistic feature of accounting for agents’ confidence [13,14]. Furthermore the variable- N model can be used to generate statistical and dynamical features similar to those observed in financial markets (archetypal examples of complex systems) [2,13]. Therefore, demonstration of predictability of extreme events in the present multiagent model would open up the exciting possibility of predictability of extreme events in real-world systems. Such predictability goes beyond the standard economic paradigm of the efficient market hypothesis [10].

The only global information available to the agents is a common bit-string “memory” of the m most recent outcomes. The agents can thus be said to exhibit “bounded rationality” [5]. Consider $m = 2$; the $2^m = 4$ possible history bit strings are 00, 01, 10, and 11. A strategy consists of a response, i.e., 0 or 1, to each possible bit string; hence there are $2^{2^m} = 16$ possible strategies. At the beginning of the game, each agent randomly picks q strategies and, after each turn, assigns one (virtual) point to a strategy which would have predicted the correct outcome. Agents have a time horizon T , over which virtual points are collected, and a threshold probability level τ ; strategies with a probability of winning greater than or equal to τ , i.e., having $\geq T\tau$ virtual points, are available to be used by the agent. We call these *active* strategies. Agents with no active strategies within their individual set of q strategies do not play at that time step. Agents with one or more active strategies play the one with the highest virtual point score; any ties between active strategies are resolved using a coin toss. The excess demand N_{0-1} , which can be identified as the output from the model system, can be expressed as

$$N_{0-1} = \sum_i n_i(1 - 2s_i), \quad (1)$$

where s_i is the prediction of the i th strategy, e.g., 0 or 1, and n_i is the number of agents using this strategy; the sum is taken over the set of active strategies at that time step.

Because of the feedback in the game, any particular strategy’s success is short lived. If all the agents begin to

use similar strategies, and hence make the same decision, such a strategy ceases to be profitable. The game can be broadly classified into three regimes: (i) The number of strategies in play is much greater than the total available: groups of traders will play using the same strategy and therefore crowds should dominate the game [15]. (ii) The number of strategies in play is much less than the total available: grouping behavior is therefore minimal. (iii) The number of strategies in play is comparable to the total number available: this represents a transition regime and is of the most interest, since it produces seemingly random dynamics with occasional large movements. Remarkably, however, we find that *large* changes over several consecutive time steps can be predicted with surprising accuracy.

Suppose we are given a time series $H(t)$ with increments $\Delta H(t)$ generated by a physical, sociological, biological, or economic system (e.g., a financial market [13]), whose dynamics are well described by the multiagent game for a fixed *unknown* parameter set m, N, τ, T and an *unknown* specific realization of initial strategy choices. We call this our “black-box” game. Even with complete knowledge of the game’s state, subsequent outcomes are not perfectly predictable since the coin tosses which resolve ties in decisions (i.e., $N_0 = N_1$) and active-strategy scores inject stochasticity into the game’s time evolution. Previous authors have demonstrated the existence of a degree of stationary predictability in the basic MG, e.g., via the histogram of bit-string occurrences [16]; our results are, by contrast, dynamic. Our goal is to identify “third-party” games which can be matched with the black-box game [$\Delta H(t)$ being proportional to the excess demand N_{0-1} , or a known nonlinear function thereof] and then used to predict large future changes in $H(t)$. For the remainder of this article, we focus on the following game parameters for the black-box game: $N = 101$, $m = 3$, $q = 2$, $T = 100$, and $\tau = 0.53$, although our conclusions are more general [17]. Since $\tau > 0.5$, an agent will not participate unless she believes she has a better than average chance of winning. Note that it is computationally impractical to have large values of m in the third-party game, because there are 2^{2^m} strategies. However, we have found that the reduced strategy space, comprising a subset of 2^{m+1} strategies which are either anticorrelated or uncorrelated with each other [11], can be used to match a black-box game which was generated using the full strategy space [17].

We start by running $H(t)$ through a trial third-party game in order to generate an estimate of S_0 and S_1 at each time step, the number of active strategies predicting a 0 or 1, respectively. This is obtained from the strategy space, or the pool of all available strategies in the third-party game, and is independent of the distribution of agents. We wish to predict $\Delta H(t)$, i.e., N_{0-1} ; we will do this by linking S and N through an appropriate probability distribution. Provided the strategy space in the black-box game is reasonably well covered by the agent’s random choice of initial strategies, any bias towards a particular outcome in

the active strategy set will propagate itself as a bias in the value of N_{0-1} away from zero. Thus we expect N_{0-1} to be approximately proportional to $S_0 - S_1 = S_{0-1}$. This is equivalent to assuming an equal weighting n_i on each strategy in Eq. (1), indicating that the exact distribution of strategies among the individual agents is unimportant in this regime [18]. In addition, the number of agents taking part in the game at each time step will be related to the total number of active strategies $S_0 + S_1 = S_{0+1}$, hence the error (i.e., variance) in the prediction of N_{0-1} using S_{0-1} will depend on S_{0+1} . Based on extensive statistical analysis of known simulations for the multiagent game [17], we have confirmed that it is reasonable to model the relationship by

$$N_{0-1} = bS_{0-1} + \varepsilon[0, f(S_{0+1})],$$

where ε is a noise term with mean zero and variance a function of S_{0+1} , and b is a constant. In particular, we describe the forecast for N_{0-1} as a normal distribution of the form $N_{0-1} \sim N(bS_{0-1}, cS_{0+1})$, where c is a constant. (We seek the simplest stable distribution as a density forecast, while acknowledging that the true distribution of N_{0-1} is indeed fat tailed.)

The variance of our forecast density function can be minimized by choosing a third-party game that achieves the maximum correlation between N_{0-1} and our explanatory variable S_{0-1} , with the unexplained variance being characterized by a linear function of S_{0+1} . We focus on the parameter regime known to produce realistic statistics (e.g., fat-tailed distribution of returns in financial markets). Within this parameter space we run an ensemble of third-party games through the black-box series $H(t)$, calculating the values of S_{0-1} from the reconstructed strategy space. We then identify the configuration that achieves the highest correlation between S_{0-1} and N_{0-1} produced by the original black-box game. As shown in Fig. 1, the third-party game that achieves the highest correlation is the one whose parameters coincide with the black-box game. From a knowledge of just $H(t)$, and hence N_{0-1} , we have therefore used next-step prediction to recover all the parameters of relevance to produce a “model” game for prediction purposes. The games reported here were all homogeneous in T and τ , but we have also carried out studies in which the values of these parameters vary between agents [17]. Even if the black-box game is heterogeneous, prediction by a homogeneous third-party game still exhibits a significant degree of correlation, indicating the robustness of our procedure.

We now extend this forecast an arbitrary number j of time steps into the future, in order to address the predictability of large changes in $H(t)$ over several consecutive time steps. This is achieved by calculating the net value of S_{0-1} along all the 2^{j-1} possible future routes of the third-party game, weighted by appropriate probabilities. In order to assign these probabilities, it is necessary to calculate all possible S_{0-1} values in the next j time steps. This is possible since the only data required to update the

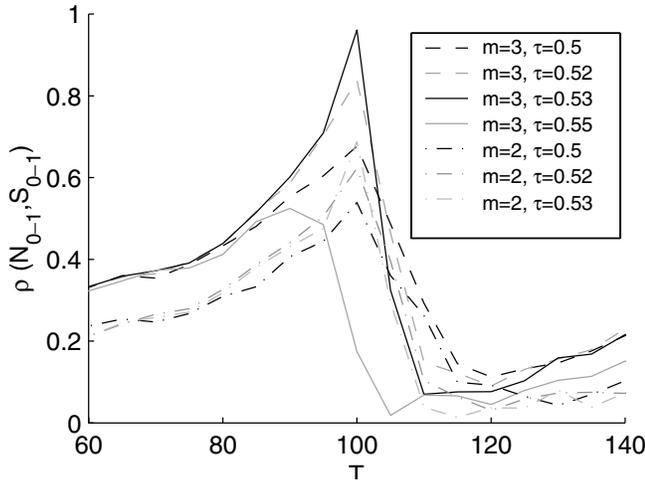


FIG. 1. Estimation of the parameter set for the black-box game. The correlation between N_{0-1} and S_{0-1} is calculated over 200 time steps for an ensemble of candidate third-party games. The third-party game that achieves the highest correlation is the one with the same parameters as the black-box game.

strategy space between time steps is knowledge of the winning decision, and hence the third-party game can be directed along a given path independent of the predictions of the individual agents in the black-box game. The change in N_{0-1} along a path indexed by k is given by a convolution of the predictions over the j individual steps and is distributed as

$$N(\mu_k, \sigma_k) \sim N\left(b \sum S_{0-1}, c \sum S_{0+1}\right),$$

where the summation is taken along the path represented by k . In general, the pdf for the change in N_{0-1} during the next j time steps is a mixture of normals:

$$P[\Delta N_{0-1}(i; i + j)] = \sum_{k=1}^{2^{j-1}} p_k N(\mu_k, \sigma_k), \quad (2)$$

where p_k is the probability of path k being taken.

To test the validity of the density forecast, we perform a statistical evaluation using the realized variables. The one-step-ahead forecasts are normal distributions, and we define the test statistic Z_i as

$$Z_i = \frac{x_i - \mu_i^x}{\sigma_i^x}, \quad (3)$$

where μ_i^x and σ_i^x are the mean and variance of the forecast distribution, and x_i is the realized value of N_{0-1} at time step i . The Z_i were found to be independent uniform $N(0, 1)$ variates for 1000 out-of-sample predictions, confirming that the predicted distributions are correct. To compare the forecasts to a naive “no-change” prediction, we calculate the Theil coefficient [19] which is the sum of squared prediction errors divided by the sum of squared errors resulting from the naive forecast. A coefficient of less than 1 implies a superior performance compared to the naive prediction; calculated values were typically in the region of 0.4. There is no accepted method in the literature

for evaluating multi-step-ahead forecasts [20]. However, the density function for an arbitrary time horizon is a mixture of normal distributions, see Eq. (2), each of which can be roughly characterized in terms of a single mean and variance:

$$E[X] = \sum_{k=1}^{2^{j-1}} p_i \mu_i,$$

$$\text{Var}[X] = \sum_{k=1}^{2^{j-1}} p_i (\sigma_i^2 + \mu_i^2) - \left(\sum_{k=1}^{2^{j-1}} p_i \mu_i \right)^2.$$

Hence the same test statistic as Eq. (3) can be calculated. Again, the predictions were found to be reliable.

Given that we can derive accurate distributions for the future changes in $H(t)$, these will be of most practical interest in situations where there is likely to be a substantial, well-defined movement. We characterize these moments by seeking distributions with a high value of $|\mu|$ and a low value of σ at a future time step, or over a specified time horizon. In Fig. 2 we plot $|\mu|$ vs σ for a number of separate forecasts, and take a fraction of points that are farthest from the average trend indicated by the regression line, i.e., we are interested in the outliers. The point with the highest residual is thus a candidate for the game to be in a highly predictable phase. We call these time periods *predictable corridors*, since comparatively tight confidence intervals can be drawn for the future evolution of the excess demand, a typical example of which is shown in Fig. 3. A standard autoregressive prediction AR(8), which is based on information from the previous eight time steps, does not pick up the large change. Furthermore, no significant

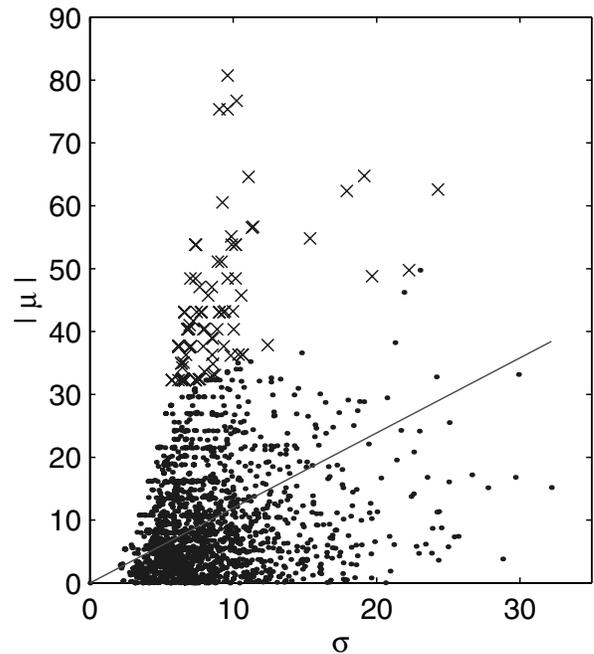


FIG. 2. A plot of $|\mu|$ vs σ for 500 separate four-step density forecasts. Items marked by “x” are forecasts with an unusually large value of $|\mu|/\sigma$. At these moments, the game is likely to be in a highly predictable phase.

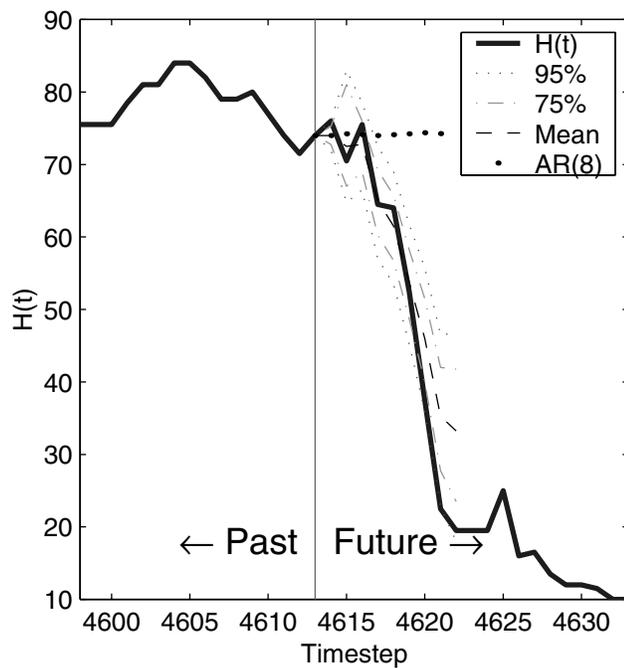


FIG. 3. Comparison between the forecast density function and the realized time series $H(t)$ for a typical large movement. The large, well-defined movement is correctly predicted. An AR(8)-based prediction has been included for comparison.

linear autocorrelation (at the 95% level) exists prior to the large movement studied. We subject these points to an identical test as described earlier to ensure that these potential outliers are well described by our probability distributions, and this is found to be true. We note that the coin-toss frequency does not change dramatically prior to the large movements, confirming our statement that the large changes are global and hence cannot be traced to a single nucleation event [17].

We performed extensive numerical simulations to check the validity of these predictive corridors [17]. Our procedure is to take a sample of 5000 time steps, then fit parameters using the first 3000 steps. We then look at the largest changes (extreme events) in our out-of-sample region. Extreme events are ranked by the largest movements in $H(t)$ over a given window size W . Hence we consider the top twenty extreme events and calculate the probability integral transform z_t of the realized variables with respect to the forecast densities. The z_t are found to be approximately uniform $U[0, 1]$ variates, confirming that the forecast distribution is essentially correct. About 50% of large movements occur in periods with tight predictable corridors, i.e., a large value of $|\mu|/\sigma$. Both the magnitude and sign of these extreme events are therefore predictable. The remainder correspond to periods with very wide corridors. Although the magnitude of the future movement is now uncertain, the present method predicts with high probability

the actual direction of change. Even this more limited information would be invaluable for assessing future risk in the physical, economic, sociological, or biological system of interest. Finally we note that some empirical support for our claim of enhanced predictability prior to extreme movements has very recently appeared for the case of financial markets [21].

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