

Limit theorems under sublinear expectations and probabilities

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Outline

- Introduction
- Law of large numbers
- Central limit theorem

Introduction

- Capacity and Choquet expectation

 Marinacci, M. (1999) Limit laws for non-additive probabilities and their frequentist interpretation. *J. Econom. Theory*, 84, 145-195.

 Maccheroni, F. and Marinacci, M. (2005) A strong law of large numbers for capacities. *The Annals of Probability*, 33, 1171-1178.

etc.

- Sublinear expectation

 Peng, S. (2008) A new central limit theorem under sublinear expectations. in [arXiv:0803.2656v1](https://arxiv.org/abs/0803.2656v1).

 Chen, Z. (2010) Strong laws of large numbers for capacities. in [arXiv:1006.0749v1](https://arxiv.org/abs/1006.0749v1).

etc.

Let (Ω, \mathcal{F}) be a measurable space and \mathcal{M} be the set of all probabilities on Ω .

For each non-empty subsets $\mathcal{P} \subset \mathcal{M}$, we can define:

- Upper probability $V(A) := \sup_{P \in \mathcal{P}} P(A)$, $A \in \mathcal{F}$
- Lower probability $v(A) := \inf_{P \in \mathcal{P}} P(A)$, $A \in \mathcal{F}$
- Sublinear expectation $\hat{\mathbb{E}}[X] := \sup_{P \in \mathcal{P}} E_P[X]$, i.e.

$$(1) \hat{\mathbb{E}}[X] \leq \hat{\mathbb{E}}[Y] \text{ if } X \leq Y$$

$$(2) \hat{\mathbb{E}}[c] = c, \forall c \in \mathbb{R}$$

$$(3) \hat{\mathbb{E}}[X + Y] \leq \hat{\mathbb{E}}[X] + \hat{\mathbb{E}}[Y]$$

$$(4) \hat{\mathbb{E}}[\lambda X] = \lambda \hat{\mathbb{E}}[X], \forall \lambda \geq 0$$

Definition (Independence)

- X_n is said to be **independent** of (X_1, \dots, X_{n-1}) , if for each $\varphi \in C_{b,Lip}(\mathbb{R}^n)$ (all bounded and Lipschitz functions on \mathbb{R}^n),

$$\hat{\mathbb{E}}[\varphi(X_1, \dots, X_n)] = \hat{\mathbb{E}}[\hat{\mathbb{E}}[\varphi(x_1, \dots, x_{n-1}, X_n)]|_{(x_1, \dots, x_{n-1})=(X_1, \dots, X_{n-1})}].$$

- X_n is said to be **product independent** of (X_1, \dots, X_{n-1}) if for each nonnegative bounded Lipschitz function φ_k ,

$$\hat{\mathbb{E}}\left[\prod_{k=1}^n \varphi_k(X_k)\right] = \prod_{k=1}^n \hat{\mathbb{E}}[\varphi_k(X_k)].$$

- X_n is said to be **sum independent** of (X_1, \dots, X_{n-1}) if for each $\varphi \in C_{b,Lip}(\mathbb{R})$,

$$\hat{\mathbb{E}}\left[\varphi\left(\sum_{k=1}^n X_k\right)\right] = \hat{\mathbb{E}}\left[\hat{\mathbb{E}}[\varphi(x + X_n)]|_{x=\sum_{k=1}^{n-1} X_k}\right].$$

Example:

We consider X and Y such that

$$-\hat{\mathbb{E}}[-X] < \hat{\mathbb{E}}[X] = 0 \quad \text{and} \quad -\hat{\mathbb{E}}[-Y] < \hat{\mathbb{E}}[Y] = 0.$$

Independent case:

$$\begin{aligned}\hat{\mathbb{E}}[XY] &= \hat{\mathbb{E}}[\hat{\mathbb{E}}[xY]|_{x=X}] \\ &= \hat{\mathbb{E}}[(x^+ \hat{\mathbb{E}}[Y] + x^- \hat{\mathbb{E}}[-Y])|_{x=X}] \\ &= \hat{\mathbb{E}}[X \hat{\mathbb{E}}[Y] + X^- (\hat{\mathbb{E}}[Y] + \hat{\mathbb{E}}[-Y])] \\ &> 0\end{aligned}$$

Definition (Identical distribution)

X_2 is said to be identically distributed with X_1 , denoted by $X_1 \sim X_2$, if for each $\varphi \in C_{b,Lip}(\mathbb{R})$,

$$\hat{\mathbb{E}}[\varphi(X_1)] = \hat{\mathbb{E}}[\varphi(X_2)].$$

Definition (G -normal distribution)

ξ is said to be G -normal distributed, denoted by $\xi \sim \mathcal{N}(0, [\underline{\sigma}^2, \bar{\sigma}^2])$, where $\bar{\sigma}^2 = \hat{\mathbb{E}}[\xi^2]$ and $\underline{\sigma}^2 = -\hat{\mathbb{E}}[-\xi^2]$, if

$$\forall a, b \geq 0, a\xi + b\bar{\xi} \sim \sqrt{a^2 + b^2}\xi,$$

where $\bar{\xi}$ is independent of ξ and $\xi \sim \bar{\xi}$.

Remark

If $\xi + \bar{\xi} \sim \sqrt{2}\xi$, then ξ is G -normal distributed.

Proposition

If ξ is G -normal distributed with $\bar{\sigma}^2 = \hat{\mathbb{E}}[\xi^2]$ and $\underline{\sigma}^2 = -\hat{\mathbb{E}}[-\xi^2]$, for each $\varphi \in C_{b,Lip}(\mathbb{R})$, we define $u(t, x) = \hat{\mathbb{E}}[\varphi(x + \sqrt{t}\xi)]$, $(t, x) \in [0, \infty) \times \mathbb{R}$, then $u(t, x)$ is the unique viscosity solution of the following G -heat PDE:

$$\partial_t u - G(\partial_{xx}^2 u) = 0, u|_{t=0} = \varphi,$$

where $G(\alpha) = \frac{1}{2}(\bar{\sigma}^2 \alpha^+ - \underline{\sigma}^2 \alpha^-)$, $\alpha \in \mathbb{R}$.

Proposition

Let $\xi \sim \mathcal{N}(0, [\underline{\sigma}^2, \bar{\sigma}^2])$, if φ is a convex function, then $\hat{\mathbb{E}}[\varphi(\xi)] = E_Z[\varphi(\bar{\sigma}Z)]$, where $Z \sim \mathcal{N}(0, 1)$.

Definition (Maximal distribution)

η is said to be maximal distributed, if

$$\forall a, b \geq 0, a\eta + b\bar{\eta} \sim (a + b)\eta.$$

Remark

If $\eta + \bar{\eta} \sim 2\eta$, then η is maximal distributed.

Proposition

If η is maximal distributed with $\bar{\mu} = \hat{\mathbb{E}}[\eta]$ and $\underline{\mu} = -\hat{\mathbb{E}}[-\eta]$, then for each

$\varphi \in C_{b,Lip}(\mathbb{R})$,

$$\hat{\mathbb{E}}[\varphi(\eta)] = \max_{\underline{\mu} \leq \mu \leq \bar{\mu}} \varphi(\mu).$$

Law of large numbers

Weak law of large numbers (Peng)

Let $\{X_n\}$ be a sequence of i.i.d random variables with finite means $\bar{\mu} = \hat{\mathbb{E}}[X_1]$ and $\underline{\mu} = -\hat{\mathbb{E}}[-X_1]$. Suppose $\hat{\mathbb{E}}[|X_1|^{1+\alpha}] < \infty$ for some $\alpha > 0$, let $S_n = \sum_{k=1}^n X_k$, then

$$\lim_{n \rightarrow \infty} \hat{\mathbb{E}}\left[\varphi\left(\frac{S_n}{n}\right)\right] = \max_{\underline{\mu} \leq \mu \leq \bar{\mu}} \varphi(\mu), \quad \forall \varphi \in C_{b,Lip}(\mathbb{R}).$$

Strong law of large numbers (Chen)

Let $\{X_n\}$ be a sequence of i.i.d random variables with finite means $\bar{\mu} = \hat{\mathbb{E}}[X_1]$ and $\underline{\mu} = -\hat{\mathbb{E}}[-X_1]$. Suppose $\hat{\mathbb{E}}[|X_1|^{1+\alpha}] < \infty$ for some $\alpha > 0$, let $S_n = \sum_{k=1}^n X_k$, then

(I) $v(\underline{\mu} \leq \liminf_{n \rightarrow \infty} \frac{S_n}{n} \leq \limsup_{n \rightarrow \infty} \frac{S_n}{n} \leq \bar{\mu}) = 1$.

(II) Furthermore, if V is upper continuous, i.e., $V(A_n) \downarrow V(A)$, if $A_n \downarrow A$, then $V(\limsup_{n \rightarrow \infty} \frac{S_n}{n} = \bar{\mu}) = 1$, $V(\liminf_{n \rightarrow \infty} \frac{S_n}{n} = \underline{\mu}) = 1$.

(III) Suppose that V is upper continuous, and $C(\{x_n\})$ is the cluster set of a sequence of $\{x_n\}$ in \mathbb{R} , i.e.,

$C(\{x_n\}) = \{x \mid \text{there exists a subsequence } \{x_{n_k}\} \text{ of } \{x_n\} \text{ such that } x_{n_k} \rightarrow x\}$, then $V(C(\{\frac{S_n}{n}\})) = [\underline{\mu}, \bar{\mu}] = 1$.

Law of large numbers

We assume \mathcal{P} is **weakly compact**. Let $\{X_k\}_{k=1}^{\infty}$ be a sequence of random variables satisfying: $\sup_{k \geq 1} \hat{\mathbb{E}}[|X_k|^{1+\alpha}] < \infty$, for some $\alpha >$

0, and $\hat{\mathbb{E}}[X_k] \equiv \bar{\mu}$, $-\hat{\mathbb{E}}[-X_k] \equiv \underline{\mu}$, $k = 1, 2, \dots$. Set $S_n = \sum_{k=1}^n X_k$.

(i) If $\{X_k\}_{k=1}^{\infty}$ is product independent, then

$$v(\underline{\mu} \leq \liminf_{n \rightarrow \infty} \frac{S_n}{n} \leq \limsup_{n \rightarrow \infty} \frac{S_n}{n} \leq \bar{\mu}) = 1.$$

(ii) If $\{X_k\}_{k=1}^{\infty}$ is product and sum independent, then

$$\lim_{n \rightarrow \infty} \hat{\mathbb{E}}[\varphi(\frac{S_n}{n})] = \max_{\underline{\mu} \leq \mu \leq \bar{\mu}} \varphi(\mu).$$

(iii) If $\{X_k\}_{k=1}^{\infty}$ is sum independent, and $V(\cdot)$ is upper continuous, then

$$\forall \mu \in [\underline{\mu}, \bar{\mu}], \quad V(\lim_{n \rightarrow \infty} \frac{S_n}{n} = \mu) = 1.$$

Lemma

We assume that \mathcal{P} is weakly compact. Let P be a probability measures such that for each $\varphi \in C_{b,Lip}(\mathbb{R}^n)$,

$$E_P[\varphi(X_1, \dots, X_n)] \leq \hat{\mathbb{E}}[\varphi(X_1, \dots, X_n)], \quad (*)$$

Then $(*)$ also holds for each bounded measurable function φ .

Proof: $(*)$ holds for bounded u.s.c. function φ since $\exists \varphi_k \in C_{b,Lip}(\mathbb{R}^n)$, s.t. $\varphi_k \downarrow \varphi$ ($\varphi_k = \sup_{y \in \mathbb{R}^n} \{\varphi(y) - k\|x - y\|\}$). \mathcal{P} is **weakly compact**, we have $\hat{\mathbb{E}}[\varphi_k(X_1, \dots, X_n)] \downarrow \hat{\mathbb{E}}[\varphi(X_1, \dots, X_n)]$ (Theorem 31 in Denis, Hu, Peng(2008)), then $E_P[\varphi(X_1, \dots, X_n)] = \lim_{k \rightarrow \infty} E_P[\varphi_k(X_1, \dots, X_n)] \leq \hat{\mathbb{E}}[\varphi(X_1, \dots, X_n)]$.

If φ is a bounded measurable function, then

$$\begin{aligned} & E_P[\varphi(X_1, \dots, X_n)] \\ &= \sup\{E_P[\bar{\varphi}(X_1, \dots, X_n)] : \bar{\varphi} \text{ is bounded upper semi-continuous and } \bar{\varphi} \leq \varphi\} \\ &\leq \hat{\mathbb{E}}[\varphi(X_1, \dots, X_n)]. \end{aligned}$$

Lemma

If for each nonnegative bounded Lipschitz function φ_i , $i = 1, \dots, n$,

$$\hat{\mathbb{E}}\left[\prod_{i=1}^n \varphi_i(X_i)\right] = \prod_{i=1}^n \hat{\mathbb{E}}[\varphi_i(X_i)],$$

then for each nonnegative bounded measurable function φ_i , $i = 1, \dots, n$,

$$\hat{\mathbb{E}}\left[\prod_{i=1}^n \varphi_i(X_i)\right] = \prod_{i=1}^n \hat{\mathbb{E}}[\varphi_i(X_i)].$$

Sketch proof of main theorem:

(i) Similar to Proof of Chen (2010)

Consider $Y_k = X_k I_{\{|X_k| \leq k^\beta\}} - \hat{\mathbb{E}}[X_k I_{\{|X_k| \leq k^\beta\}}]$, where $\frac{1}{1+\alpha} < \beta < 1$.

$$e^x \leq 1 + x + |x|^{1+\alpha} e^{2|x|}, x \in \mathbb{R}, 0 < \alpha < 1.$$

$$e^{\frac{Y_k}{n^\beta}} \leq 1 + \frac{Y_k}{n^\beta} + \frac{|Y_k|^{1+\alpha}}{n^{\beta(1+\alpha)}} e^{2|\frac{Y_k}{n^\beta}|} \leq 1 + \frac{Y_k}{n^\beta} + \frac{|Y_k|^{1+\alpha}}{n} e^4.$$

$$\hat{\mathbb{E}}[e^{\frac{Y_k}{n^\beta}}] \leq 1 + \frac{C}{n}.$$

$$\hat{\mathbb{E}}\left[\prod_{k=1}^n e^{\frac{Y_k}{n^\beta}}\right] = \prod_{k=1}^n \hat{\mathbb{E}}[e^{\frac{Y_k}{n^\beta}}] \leq \left(1 + \frac{C}{n}\right)^n \leq e^C.$$

$$V\left(\frac{\sum_{k=1}^n Y_k}{n} \geq \varepsilon\right) = V\left(e^{\frac{\sum_{k=1}^n Y_k}{n^\beta}} \geq e^{\varepsilon n^{1-\beta}}\right) \leq e^{-\varepsilon n^{1-\beta}} \hat{\mathbb{E}}\left[\prod_{k=1}^n e^{\frac{Y_k}{n^\beta}}\right] \leq e^{-\varepsilon n^{1-\beta}} e^C.$$

Corollary

Let $\{X_n\}$ be a sequence of bounded and sum independent random variables satisfying: $\sup_{k \geq 1} \hat{\mathbb{E}}[|X_k|^{1+\alpha}] < \infty$, for some $\alpha > 0$, and $\hat{\mathbb{E}}[X_k] \equiv \bar{\mu}$, $-\hat{\mathbb{E}}[-X_k] \equiv \underline{\mu}$, $k = 1, 2, \dots$. Set $S_n = \sum_{k=1}^n X_k$.

(i) $v(\underline{\mu} \leq \liminf_{n \rightarrow \infty} \frac{S_n}{n} \leq \limsup_{n \rightarrow \infty} \frac{S_n}{n} \leq \bar{\mu}) = 1$.

(ii) $\lim_{n \rightarrow \infty} \hat{\mathbb{E}}[\varphi(\frac{S_n}{n})] = \max_{\underline{\mu} \leq \mu \leq \bar{\mu}} \varphi(\mu)$.

(iii) Furthermore, if V is upper continuous, then

$$\forall \mu \in [\underline{\mu}, \bar{\mu}], \quad V\left(\lim_{n \rightarrow \infty} \frac{S_n}{n} = \mu\right) = 1.$$

Central limit theorem

Central limit theorem (Peng)

Let $\{X_n\}$ be a sequence of i.i.d. random variables such that $\hat{\mathbb{E}}[X_1] = \hat{\mathbb{E}}[-X_1] = 0$ and $\hat{\mathbb{E}}[|X_1|^q] < \infty$ for some $q > 2$. Let $S_n = \sum_{k=1}^n X_k$. Then for all $\varphi \in C(\mathbb{R})$ with quadratic growth condition,

$$\lim_{n \rightarrow \infty} \hat{\mathbb{E}}\left[\varphi\left(\frac{S_n}{\sqrt{n}}\right)\right] = \hat{\mathbb{E}}[\varphi(\xi)],$$

where $\xi \sim \mathcal{N}(0, [\underline{\sigma}^2, \bar{\sigma}^2])$ with $\bar{\sigma}^2 = \hat{\mathbb{E}}[X_1^2]$, $\underline{\sigma}^2 = -\hat{\mathbb{E}}[-X_1^2]$.

Let $\mathcal{M}_n^q([\underline{\sigma}^2, \bar{\sigma}^2], K)$ denote the set of n -stages martingale S with filtration \mathcal{F} , such that for all k , both relations hold:

- $E[S_k] = 0$
- $\underline{\sigma}^2 \leq E[|S_{k+1} - S_k|^2 | \mathcal{F}_k] \leq \bar{\sigma}^2$
- $E[|S_{k+1} - S_k|^q | \mathcal{F}_k] \leq K^q$

Let

$$V_n[\varphi] := \sup_{S \in \mathcal{M}_n^q([\underline{\sigma}^2, \bar{\sigma}^2], K)} E\left[\varphi\left(\frac{S_n}{\sqrt{n}}\right)\right].$$

Theorem (Central limit theorem)

We assume that $q > 2$. Then for all $\varphi \in C(\mathbb{R})$ with quadratic growth condition, we have

$$\lim_{n \rightarrow \infty} V_n[\varphi] = \hat{\mathbb{E}}[\varphi(\xi)],$$

where $\xi \sim \mathcal{N}(0, [\underline{\sigma}^2, \bar{\sigma}^2])$.

Maximal L^p variation problem

Let $\mathcal{M}_n(\mu)$ be the set of n -stage martingales whose terminal distribution is μ . We define the L^p -variation of length n of the martingale $(L_k)_{k=1, \dots, n}$ as

$$\mathcal{V}_n^p(L) = E\left[\sum_{k=1}^n (E[|L_k - L_{k-1}|^p | (L_i, i \leq k-1)])^{\frac{1}{p}}\right].$$

The value function is denoted by






$$V_n(\mu) = \frac{1}{\sqrt{n}} \sup_{L \in \mathcal{M}_n(\mu)} \mathcal{V}_n^p(L).$$

Theorem

For $p \in [1, 2)$ and $\mu \in \Delta(K)$, where K is a compact subset of \mathbb{R} . We have

$$\lim_{n \rightarrow \infty} V_n(\mu) = E[f_\mu(Z)Z],$$

where $Z \sim N(0, 1)$ and $f_\mu : \mathbb{R} \rightarrow \mathbb{R}$ is an increasing function such that $f_\mu(Z) \sim \mu$, i.e., $f_\mu(x) = F_\mu^{-1}(F_N(x))$ with $F_\mu^{-1} = \inf\{s : F_\mu(s) > y\}$.

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Sketch Proof of Theorem:

$$\sup_{L \sim \mu, S \in \mathcal{M}'_n} E\left[L \frac{S_n}{\sqrt{n}}\right] \leq V_n(\mu) \leq \sup_{L \sim \mu, S \in \mathcal{M}_n^q([0,1],2)} E\left[L \frac{S_n}{\sqrt{n}}\right].$$

$$\sup_{L \sim \mu, S_n} E\left[L \frac{S_n}{\sqrt{n}}\right] = \inf_{\varphi \in \text{Conv}(K)} E\left[\varphi(L) + \varphi^*\left(\frac{S_n}{\sqrt{n}}\right)\right],$$

$$\begin{aligned} \sup_{L \sim \mu, S \in \mathcal{M}_n^q([0,1],2)} E\left[L \frac{S_n}{\sqrt{n}}\right] &= \sup_{S \in \mathcal{M}_n^q([0,1],2)} \inf_{\varphi \in \text{Conv}(K)} E\left[\varphi(L) + \varphi^*\left(\frac{S_n}{\sqrt{n}}\right)\right] \\ &= \inf_{\varphi \in \text{Conv}(K)} E\left[\varphi(L) + \sup_{S \in \mathcal{M}_n^q([0,1],2)} E\left[\varphi^*\left(\frac{S_n}{\sqrt{n}}\right)\right]\right]. \end{aligned}$$

Since φ^* is a convex function, we have

$$\lim_{n \rightarrow \infty} \sup_{S \in \mathcal{M}_n^q([0,1],2)} E[\varphi^*\left(\frac{S_n}{\sqrt{n}}\right)] = E[\varphi^*(Z)], \quad Z \sim N(0,1).$$

$$\begin{aligned} \lim_{n \rightarrow \infty} V_n(\mu) &= \inf_{\varphi \in \text{Conv}(K)} E[\varphi(L) + \lim_{n \rightarrow \infty} \sup_{S \in \mathcal{M}_n^q([0,1],2)} E[\varphi^*\left(\frac{S_n}{\sqrt{n}}\right)]] \\ &= \inf_{\varphi \in \text{Conv}(K)} E[\varphi(L) + \varphi^*(Z)] = E[f_\mu(Z)Z]. \end{aligned}$$

Thank you for your attention!