

Lecture 2: Introduction to Monte Carlo methods

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Objective

There are many situations in which we want to estimate the average value of some random quantity

In general, we

- start with a random sample ω (which might correspond to a set of random numbers)
- usually compute some intermediate quantity U
- then evaluate a scalar output $f(U)$

$$\omega \rightarrow U \rightarrow f(U)$$

The objective is then to compute the **expected** (or average) value

$$\mathbb{E}[f(U)]$$

Basics

In some cases, the random inputs are discrete: X has value x_i with probability p_i , and then

$$\mathbb{E}[f(X)] = \sum_i f(x_i) p_i$$

In other cases, the random inputs are continuous random variables: X has probability density $p(x)$ if $\mathbb{P}(X \in (x, x+dx)) \approx p(x) dx$ and then

$$\mathbb{E}[f(X)] = \int f(x) p(x) dx$$

In either case, if a, b are random variables, and λ, μ are constants,

$$\mathbb{E}[a + \mu] = \mathbb{E}[a] + \mu$$

$$\mathbb{E}[\lambda a] = \lambda \mathbb{E}[a]$$

$$\mathbb{E}[a + b] = \mathbb{E}[a] + \mathbb{E}[b]$$

Basics

The **variance** is defined as

$$\begin{aligned}\mathbb{V}[a] &\equiv \mathbb{E} \left[(a - \mathbb{E}[a])^2 \right] \\ &= \mathbb{E} \left[a^2 - 2a \mathbb{E}[a] + (\mathbb{E}[a])^2 \right] \\ &= \mathbb{E} [a^2] - (\mathbb{E}[a])^2\end{aligned}$$

It then follows that

$$\begin{aligned}\mathbb{V}[a + \mu] &= \mathbb{V}[a] \\ \mathbb{V}[\lambda a] &= \lambda^2 \mathbb{V}[a] \\ \mathbb{V}[a + b] &= \mathbb{V}[a] + 2 \text{Cov}[a, b] + \mathbb{V}[b]\end{aligned}$$

where

$$\text{Cov}[a, b] \equiv \mathbb{E} \left[(a - \mathbb{E}[a]) (b - \mathbb{E}[b]) \right]$$

Basics

X_1 and X_2 are independent continuous random variables if

$$p_{\text{joint}}(x_1, x_2) = p_1(x_1) p_2(x_2)$$

We then get

$$\begin{aligned}\mathbb{E}[f_1(X_1) f_2(X_2)] &= \iint f_1(x_1) f_2(x_2) p_{\text{joint}}(x_1, x_2) dx_1 dx_2 \\ &= \iint f_1(x_1) f_2(x_2) p_1(x_1) p_2(x_2) dx_1 dx_2 \\ &= \left(\int f_1(x_1) p_1(x_1) dx_1 \right) \left(\int f_2(x_2) p_2(x_2) dx_2 \right) \\ &= \mathbb{E}[f_1(X_1)] \mathbb{E}[f_2(X_2)]\end{aligned}$$

So, if a, b are independent, $\text{Cov}[a, b] = 0 \implies \mathbb{V}[a+b] = \mathbb{V}[a] + \mathbb{V}[b]$

More generally, the variance of the sum of independent r.v.'s is the sum of their variances.

Random Number Generation

Monte Carlo simulation starts with random number generation, usually split into 2 stages:

- generation of independent uniform $(0, 1)$ random variables
- conversion into random variables with a particular distribution (e.g. Normal)

Very important: never write your own generator, always use a well validated generator from a reputable source

- python
- MATLAB
- Intel MKL (Math Kernel Library)

Uniform Random Variables

Pseudo-random generators use a deterministic (i.e. repeatable) algorithm to generate a sequence of (apparently) random numbers on $(0, 1)$ interval.

What defines a good generator?

- a long period – how long it takes before the sequence repeats itself 2^{32} is not enough (need at least 2^{40})
- various statistical tests to measure “randomness” – well validated software will have gone through these checks

For information see

- Intel MKL information
www.intel.com/cd/software/products/asm-na/eng/266864.htm
- Matlab information
www.mathworks.com/moler/random.pdf
- Wikipedia information
en.wikipedia.org/wiki/Random_number_generation

Normal Random Variables

$N(0, 1)$ Normal random variables (mean 0, variance 1) have the probability density function

$$p(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x^2\right) \equiv \phi(x)$$

If X is a $N(0, 1)$ Normal random variable, then its CDF (Cumulative Distribution Function) is defined as

$$\mathbb{P}[X < x] = \int_{-\infty}^x \phi(x) dx \equiv \Phi(x)$$

Many maths software libraries include the function $\Phi(x)$, along with \sin , \cos , \exp , \log and others. In python it is `norm.cdf` from `scipy.stats`

Normal Random Variables

The Box-Muller transformation method takes two independent uniform $(0, 1)$ random numbers y_1, y_2 , and defines

$$\begin{aligned}x_1 &= \sqrt{-2 \log(y_1)} \cos(2\pi y_2) \\x_2 &= \sqrt{-2 \log(y_1)} \sin(2\pi y_2)\end{aligned}$$

It can be proved that x_1 and x_2 are $N(0, 1)$ random variables, and independent:

$$p_{\text{joint}}(x_1, x_2) = p(x_1) p(x_2)$$

Inverse CDF

An alternative uses the cumulative distribution function $\Phi(x)$.

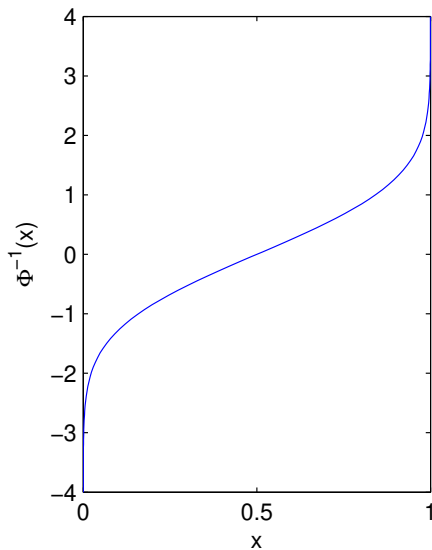
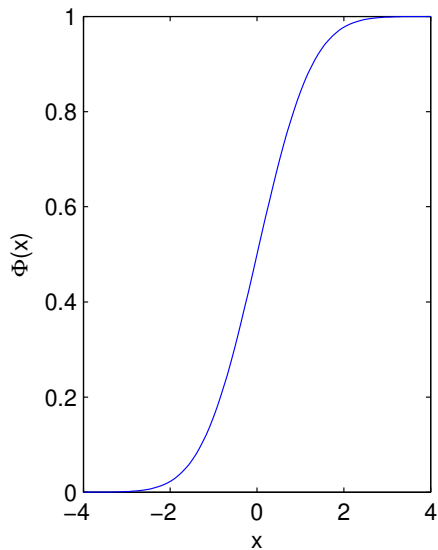
If X is a $N(0, 1)$ random variable, then $Y = \Phi(X)$ is a uniform $(0, 1)$ random variable.

Hence, can start with a uniform $(0, 1)$ random variable Y and define X by

$$X = \Phi^{-1}(Y)$$

$\Phi^{-1}(y)$ is approximated in software in a very similar way to other functions like \cos , \sin , \log , \exp . In python it is `norm.ppf` from `scipy.stats`

Normal Random Variables



Monte Carlo estimate

If we have a sequence f_n of N independent samples of f , the average

$$\bar{f} = N^{-1} \sum_{n=1}^N f_n.$$

is the Monte Carlo estimate of the expected value $\mathbb{E}[f]$

It is an unbiased estimate, since for each n ,

$$\mathbb{E}[f_n] = \mathbb{E}[f] \quad \implies \quad \mathbb{E}[\bar{f}] = \mathbb{E}[f]$$

We also have

$$\mathbb{V}[\bar{f}] = N^{-2} \mathbb{V} \left[\sum_{n=1}^N f_n \right] = N^{-2} \sum_{n=1}^N \mathbb{V}[f_n] = N^{-1} \mathbb{V}[f]$$

Central Limit Theorem

The Central Limit Theorem says that if the variance $\sigma^2 \equiv \mathbb{V}[f]$ is finite, then the error

$$e_N(f) = \bar{f} - \mathbb{E}[f]$$

is approximately Normal in distribution for large N , i.e.

$$e_N(f) \sim \sigma N^{-1/2} Z$$

where Z is a $N(0, 1)$ random variable

Central Limit Theorem

If Z is a $N(0, 1)$ random variable, then due to symmetry

$$\mathbb{P}[|Z| < s] = 1 - \mathbb{P}[|Z| > s] = 1 - 2\mathbb{P}[Z < -s] = 1 - 2\Phi(-s)$$

Table of probabilities for different s :

s	1.0	2.0	3.0	4.0
Prob	0.683	0.9545	0.9973	0.99994

Hence, with probability 99.7%, $|e_N(f)| < 3\sigma N^{-1/2}$

$$\implies \mathbb{E}[f] \in (\bar{f} - 3\sigma N^{-1/2}, \bar{f} + 3\sigma N^{-1/2})$$

This is the confidence interval for $\mathbb{E}[f]$

Estimated Variance

Given N samples, the empirical variance is

$$\tilde{\sigma}^2 = N^{-1} \sum_{n=1}^N (f_n - \bar{f})^2 = \bar{f}^2 - (\bar{f})^2$$

where

$$\bar{f} = N^{-1} \sum_{n=1}^N f_n, \quad \bar{f}^2 = N^{-1} \sum_{n=1}^N f_n^2$$

$\tilde{\sigma}^2$ is a slightly biased estimator for σ^2 ; an unbiased estimator is

$$\hat{\sigma}^2 = (N-1)^{-1} \sum_{n=1}^N (f_n - \bar{f})^2 = \frac{N}{N-1} \left(\bar{f}^2 - (\bar{f})^2 \right)$$

Finance Applications

Geometric Brownian motion for a single asset:

$$S_T = S_0 \exp \left(\left(r - \frac{1}{2} \sigma^2 \right) T + \sigma W_T \right)$$

W_T is $N(0, T)$ random variable, so can put

$$W_T = \sqrt{T} Z$$

where Z is a $N(0, 1)$ random variable.

We are then interested in the price of financial options which can be expressed as

$$V = \mathbb{E} [f(S_T)]$$

for some “payoff” function $f(S)$.

Finance Applications

For the European call option,

$$f(S) = \exp(-rT) \max(S - K, 0)$$

while for the European put option

$$f(S) = \exp(-rT) \max(K - S, 0)$$

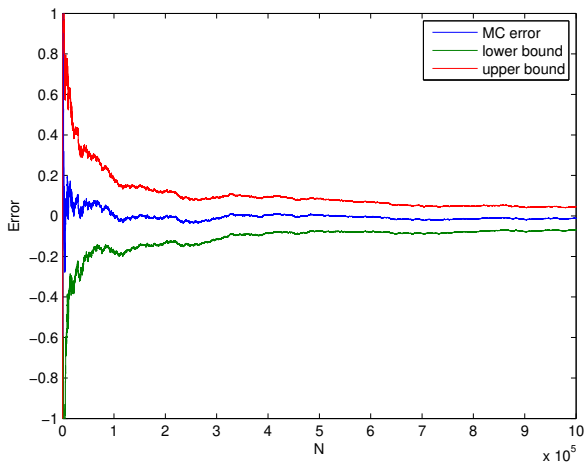
where K is the strike price

For numerical experiments we will consider a European call with
 $r=0.05$, $\sigma = 0.2$, $T=1$, $S_0=110$, $K=100$.

The analytic value is known for comparison.

Finance Applications

MC calculation with up to 10^6 paths; true value = 17.663



Finance Applications

The upper and lower bounds are given by

$$\text{Mean} \pm \frac{3 \tilde{\sigma}}{\sqrt{N}},$$

so more than a 99.7% probability that the true value lies within these bounds.

SDEs in Finance

In computational finance, stochastic differential equations are used to model the behaviour of

- stocks
- interest rates
- exchange rates
- weather
- electricity/gas demand
- crude oil prices
- ...

SDEs in Finance

Stochastic differential equations are just ordinary differential equations plus an additional random source term.

The stochastic term accounts for the uncertainty of unpredictable day-to-day events.

The aim is **not** to predict exactly what will happen in the future, but to predict the probability of a range of possible things that **might** happen, and compute some averages, or the probability of an excessive loss.

This is really what is known more generally as Uncertainty Quantification – the finance industry has been doing it for a long time because they have so much uncertainty.

SDEs in Finance

Examples:

- Geometric Brownian motion (Black-Scholes model for stock prices)

$$dS = r S dt + \sigma S dW$$

- Cox-Ingersoll-Ross model (interest rates)

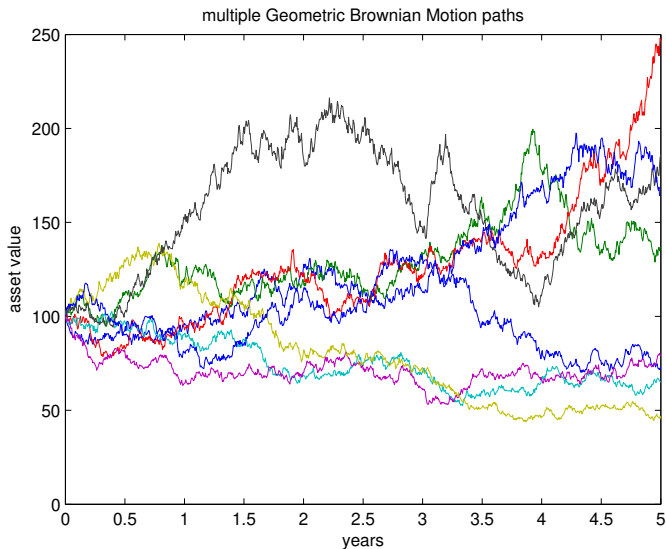
$$dr = \alpha(b - r) dt + \sigma \sqrt{r} dW$$

- Heston stochastic volatility model (stock prices)

$$\begin{aligned}dS &= r S dt + \sqrt{V} S dW_1 \\dV &= \lambda(\sigma^2 - V) dt + \xi \sqrt{V} dW_2\end{aligned}$$

with correlation ρ between dW_1 and dW_2

SDEs in Finance



Generic Problem

Stochastic differential equation with general drift and volatility terms:

$$dS_t = a(S_t, t) dt + b(S_t, t) dW_t$$

W_t is a Wiener variable with the properties that for any $q < r < s < t$, $W_t - W_s$ is Normally distributed with mean 0 and variance $t - s$, independent of $W_r - W_q$.

In many finance applications, we want to compute the expected value of an option dependent on the terminal state $P(S_T)$

Other options depend on the average, minimum and/or maximum over the whole time interval.

Euler discretisation

Given the generic SDE:

$$dS_t = a(S_t) dt + b(S_t) dW_t, \quad 0 < t < T,$$

the Euler discretisation with timestep Δt is:

$$\widehat{S}_{n+1} = \widehat{S}_n + a(\widehat{S}_n) \Delta t + b(\widehat{S}_n) \Delta W_n$$

where ΔW_n are independent Normal random variables with mean 0, variance Δt .

This will be our second model application in the course.