

# Multilevel Monte Carlo methods

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# Objectives

In presenting the multilevel Monte Carlo method, I hope to emphasise:

- the simplicity of the idea
- its flexibility
- that it's not prescriptive, more an approach
- scope for improved performance through being creative
- lots of people working on a variety of applications

I will focus on ideas rather than lots of numerical results.

## Control variate

Classic approach to variance reduction: approximate  $\mathbb{E}[f]$  using

$$N^{-1} \sum_{n=1}^N \left\{ f^{(n)} - \lambda \left( g^{(n)} - \mathbb{E}[g] \right) \right\}$$

where

- control variate  $g$  has known expectation  $\mathbb{E}[g]$
- $g$  is well correlated with  $f$ , and optimal value for  $\lambda$  can be estimated by a few samples

## Two-level Monte Carlo

If we want to estimate  $\mathbb{E}[f_1]$  but it is much cheaper to simulate  $f_0 \approx f_1$ , then since

$$\mathbb{E}[f_1] = \mathbb{E}[f_0] + \mathbb{E}[f_1 - f_0]$$

we can use the estimator

$$N_0^{-1} \sum_{n=1}^{N_0} f_0^{(n)} + N_1^{-1} \sum_{n=1}^{N_1} \left( f_1^{(n)} - f_0^{(n)} \right)$$

Two differences from standard control variate method:

- $\mathbb{E}[f_0]$  is not known, so has to be estimated
- $\lambda = 1$

## Two-level Monte Carlo

If we define

- $C_0, V_0$  to be cost and variance of  $f_0$
- $C_1, V_1$  to be cost and variance of  $f_1 - f_0$

then the total cost is

$$N_0 C_0 + N_1 C_1$$

and the variance (assuming independent estimators) is

$$N_0^{-1} V_0 + N_1^{-1} V_1$$

so for a fixed variance the cost is minimised by choosing

$$\frac{N_1}{N_0} = \frac{\sqrt{V_1/C_1}}{\sqrt{V_0/C_0}}$$

## Trivial example

- $f_1$  comes from double precision calculation
- $f_0$  comes from single precision calculation (often twice as fast on latest CPUs/GPUs)
- use the same random number generator for both calculations
  
- estimating  $V_0$  and  $V_1$  will give an optimal allocation of computational effort between single precision and double precision computations

# Multilevel Monte Carlo

Natural generalisation: given a sequence  $f_0, f_1, \dots, f_L$

$$\mathbb{E}[f_L] = \mathbb{E}[f_0] + \sum_{\ell=1}^L \mathbb{E}[f_\ell - f_{\ell-1}]$$

we can use the estimator

$$N_0^{-1} \sum_{n=1}^{N_0} f_0^{(n)} + \sum_{\ell=1}^L \left\{ N_\ell^{-1} \sum_{n=1}^{N_\ell} \left( f_\ell^{(n)} - f_{\ell-1}^{(n)} \right) \right\}$$

with independent estimation for each level

# Multilevel Monte Carlo

If we define

- $C_0, V_0$  to be cost and variance of  $f_0$
- $C_\ell, V_\ell$  to be cost and variance of  $f_\ell - f_{\ell-1}$

then the total cost is  $\sum_{\ell=0}^L N_\ell C_\ell$  and the variance is  $\sum_{\ell=0}^L N_\ell^{-1} V_\ell$ .

Using a Lagrange multiplier  $\mu^2$  to minimise the cost for a fixed variance

$$\frac{\partial}{\partial N_\ell} \sum_{k=0}^L (N_k C_k + \mu^2 N_k^{-1} V_k) = 0$$

gives

$$N_\ell = \mu \sqrt{V_\ell / C_\ell} \quad \implies \quad N_\ell C_\ell = \mu \sqrt{V_\ell C_\ell}$$

# Multilevel Monte Carlo

Setting the total variance equal to  $\varepsilon^2$  gives

$$\mu = \varepsilon^{-2} \left( \sum_{\ell=0}^L \sqrt{V_\ell C_\ell} \right)$$

and hence, the total cost is

$$\sum_{\ell=0}^L N_\ell C_\ell = \varepsilon^{-2} \left( \sum_{\ell=0}^L \sqrt{V_\ell C_\ell} \right)^2$$

in contrast to the standard cost which is approximately  $\varepsilon^{-2} V_0 C_L$ .

The MLMC cost savings are therefore:

- $V_L/V_0$ , if  $\sqrt{V_\ell C_\ell}$  increases with level
- $C_0/C_L$ , if  $\sqrt{V_\ell C_\ell}$  decreases with level

# Parametric Integration

Stefan Heinrich introduced multilevel ideas in 1999 for parametric integration, in which  $x$  is a finite-dimensional random variable, and want to estimate  $\mathbb{E}[f(x, \lambda)]$  for a range of values of the parameter  $\lambda$ .

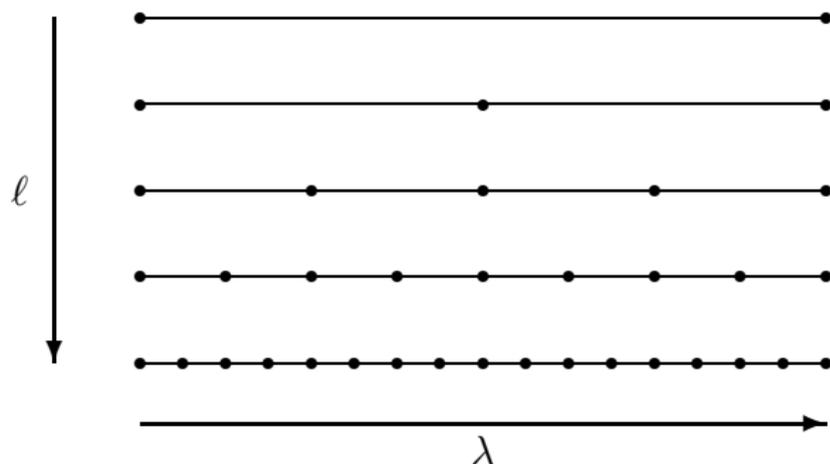
In the simplest case, suppose  $\lambda$  is a scalar, and the parameter range is  $0 \leq \lambda \leq 1$ .

If we have already estimated  $\mathbb{E}[f(x, 0)]$  and  $\mathbb{E}[f(x, 1)]$  then

$$\begin{aligned}\mathbb{E}[f(x, \frac{1}{2})] &= \frac{1}{2} \left( \mathbb{E}[f(x, 0)] + \mathbb{E}[f(x, 1)] \right) \\ &\quad + \mathbb{E} \left[ f(x, \frac{1}{2}) - \frac{1}{2}(f(x, 0) + f(x, 1)) \right]\end{aligned}$$

## Parametric Integration

This can be repeated on multiple levels (perhaps using higher order interpolation if  $f(x, \lambda)$  is sufficiently smooth)



This doesn't quite fit into the multilevel framework I've described, but the complexity analysis is very similar.

# Multilevel Path Simulation

In 2006, I introduced the multilevel approach for infinite-dimensional integration arising from SDEs driven by Brownian diffusion.

Level  $\ell$  corresponds to approximation using  $2^\ell$  timesteps, giving approximate payoff  $\widehat{P}_\ell$ .

Choice of finest level  $L$  depends on weak error (bias).

Multilevel decomposition gives

$$\mathbb{E}[\widehat{P}_L] = \mathbb{E}[\widehat{P}_0] + \sum_{\ell=1}^L \mathbb{E}[\widehat{P}_\ell - \widehat{P}_{\ell-1}]$$

# Multilevel Monte Carlo

Simplest estimator for  $\mathbb{E}[\hat{P}_\ell - \hat{P}_{\ell-1}]$  for  $\ell > 0$  is

$$\hat{Y}_\ell = N_\ell^{-1} \sum_{n=1}^{N_\ell} \left( \hat{P}_\ell^{(n)} - \hat{P}_{\ell-1}^{(n)} \right)$$

using same driving Brownian path for both levels

Variance is  $N_\ell^{-1} V_\ell$  where  $V_\ell = \mathbb{V}[\hat{P}_\ell - \hat{P}_{\ell-1}]$  gets smaller as  $\ell$  increases because  $\hat{P}_\ell, \hat{P}_{\ell-1}$  both approximate same  $P$

To make RMS error less than  $\varepsilon$

- choose  $L$  so that  $\left( \mathbb{E}[\hat{P}_L] - \mathbb{E}[P] \right)^2 < \frac{1}{2} \varepsilon^2$
- choose  $N_\ell \propto \sqrt{V_\ell / C_\ell}$  so total variance is less than  $\frac{1}{2} \varepsilon^2$

# MLMC Theorem

(Slight generalisation of original version)

If there exist independent estimators  $\widehat{Y}_\ell$  based on  $N_\ell$  Monte Carlo samples, each costing  $C_\ell$ , and positive constants  $\alpha, \beta, \gamma, c_1, c_2, c_3$  such that  $\alpha \geq \frac{1}{2} \min(\beta, \gamma)$  and

$$\text{i) } \left| \mathbb{E}[\widehat{P}_\ell - P] \right| \leq c_1 2^{-\alpha \ell}$$

$$\text{ii) } \mathbb{E}[\widehat{Y}_\ell] = \begin{cases} \mathbb{E}[\widehat{P}_0], & \ell = 0 \\ \mathbb{E}[\widehat{P}_\ell - \widehat{P}_{\ell-1}], & \ell > 0 \end{cases}$$

$$\text{iii) } \mathbb{V}[\widehat{Y}_\ell] \leq c_2 N_\ell^{-1} 2^{-\beta \ell}$$

$$\text{iv) } \mathbb{E}[C_\ell] \leq c_3 2^{\gamma \ell}$$

# MLMC Theorem

then there exists a positive constant  $c_4$  such that for any  $\varepsilon < 1$  there exist  $L$  and  $N_\ell$  for which the multilevel estimator

$$\hat{Y} = \sum_{\ell=0}^L \hat{Y}_\ell,$$

has a mean-square-error with bound  $\mathbb{E} \left[ \left( \hat{Y} - \mathbb{E}[P] \right)^2 \right] < \varepsilon^2$

with an expected computational cost  $C$  with bound

$$C \leq \begin{cases} c_4 \varepsilon^{-2}, & \beta > \gamma, \\ c_4 \varepsilon^{-2} (\log \varepsilon)^2, & \beta = \gamma, \\ c_4 \varepsilon^{-2 - (\gamma - \beta)/\alpha}, & 0 < \beta < \gamma. \end{cases}$$

# MLMC Theorem

Two observations of optimality:

- MC simulation needs  $O(\varepsilon^{-2})$  samples to achieve RMS accuracy  $\varepsilon$ .  
When  $\beta > \gamma$ , the cost is optimal —  $O(1)$  cost per sample on average.  
(Would need multilevel QMC to further reduce costs)
- When  $\beta < \gamma$ , another interesting case is when  $\beta = 2\alpha$ , which corresponds to  $\mathbb{E}[\widehat{Y}_\ell]$  and  $\sqrt{\mathbb{E}[\widehat{Y}_\ell^2]}$  being of the same order as  $\ell \rightarrow \infty$ .  
In this case, the total cost is  $O(\varepsilon^{-\gamma/\alpha})$ , which is the cost of a single sample on the finest level — again optimal.

## MLMC Theorem

MLMC Theorem allows a lot of freedom in constructing the multilevel estimator. I sometimes use different approximations on the coarse and fine levels:

$$\hat{Y}_\ell = N_\ell^{-1} \sum_{n=1}^{N_\ell} \left( \hat{P}_\ell^f(\omega^{(n)}) - \hat{P}_{\ell-1}^c(\omega^{(n)}) \right)$$

The telescoping sum still works provided

$$\mathbb{E} \left[ \hat{P}_\ell^f \right] = \mathbb{E} \left[ \hat{P}_\ell^c \right].$$

Given this constraint, can be creative to reduce the variance

$$\mathbb{V} \left[ \hat{P}_\ell^f - \hat{P}_{\ell-1}^c \right].$$

# MLMC Theorem

Two examples:

- zero-mean control variate estimator: if

$$\widehat{P}_\ell(\omega^{(n)}) \approx \widehat{P}_{\ell-1}(\omega^{(n)}) + Z(\omega^{(n)})$$

where  $\mathbb{E}[Z] = 0$ , then use

$$\widehat{P}_{\ell-1}^c(\omega^{(n)}) \equiv \widehat{P}_{\ell-1}(\omega^{(n)}), \quad \widehat{P}_\ell^f(\omega^{(n)}) \equiv \widehat{P}_\ell(\omega^{(n)}) - Z(\omega^{(n)})$$

- antithetic estimator:

$$\widehat{P}_{\ell-1}^c(\omega^{(n)}) \equiv \widehat{P}_{\ell-1}(\omega^{(n)}), \quad \widehat{P}_\ell^f(\omega^{(n)}) \equiv \frac{1}{2} \left( \widehat{P}_\ell(\omega^{(n)}) + \widehat{P}_\ell(\omega_{anti}^{(n)}) \right)$$

where  $\omega_{anti}^{(n)}$  is an antithetic “twin” with the same distribution as  $\omega^{(n)}$ .

# MLMC Challenges

- not always obvious how to couple coarse and fine levels  
i.e. what does  $\widehat{P}_\ell(\omega^{(n)}) - \widehat{P}_{\ell-1}(\omega^{(n)})$  mean?
- can the MLMC flexibility be exploited to improve the variance decay?  
— particularly important for discontinuous “payoffs”, since a small difference in the coarse and fine “paths” can produce an  $O(1)$  difference in the “payoff”
- numerical analysis – proving the rate at which  $V_\ell$  decays can be tough

# Brownian Diffusion SDEs

Brownian increments for coarse path obtained by summing increments for fine path – very simple and natural

I prefer to use the first order Milstein discretisation – for simple put / call options (and more generally Lipschitz functions of the final state) this leads to

$$\widehat{P}_\ell - \widehat{P}_{\ell-1} = O(h_\ell)$$

and hence  $V_\ell = O(h_\ell^2)$ .

However, not so easy for lookback, digital and barrier options.

(And in multiple dimensions requires Lévy areas, but can be avoided by an antithetic treatment, G & Szpruch, 2013)

## Digital options

In a digital option, the payoff is a discontinuous function of the final state.

Using the Milstein approximation, first order strong convergence means that  $O(h_\ell)$  of the simulations have coarse and fine paths on opposite sides of a discontinuity.

Hence,

$$\widehat{P}_\ell - \widehat{P}_{\ell-1} = \begin{cases} O(1), & \text{with probability } O(h_\ell) \\ O(h_\ell), & \text{with probability } O(1) \end{cases}$$

so

$$\mathbb{E}[\widehat{P}_\ell - \widehat{P}_{\ell-1}] = O(h_\ell), \quad \mathbb{E}[(\widehat{P}_\ell - \widehat{P}_{\ell-1})^2] = O(h_\ell),$$

and hence  $V_\ell = O(h_\ell)$ , not  $O(h_\ell^2)$

# Digital options

Three fixes:

- Splitting: split each path simulation into  $M$  paths by trying  $M$  different values for the Brownian increment for the last fine path timestep
- Conditional expectation: using the Euler discretisation instead of Milstein for the final timestep, conditional on all but the final Brownian increment, the final state has a Gaussian distribution, with a known analytic conditional expectation in simple cases
- Change of measure: when the expectation is not known, can use a change of measure so the coarse path takes the same final state as the fine path — difference in the “payoff” now comes from the Radon-Nikodym derivative

# Numerical Analysis

option	Euler		Milstein	
	numerics	analysis	numerics	analysis
Lipschitz	$O(h)$	$O(h)$	$O(h^2)$	$O(h^2)$
Asian	$O(h)$	$O(h)$	$O(h^2)$	$O(h^2)$
lookback	$O(h)$	$O(h)$	$O(h^2)$	$o(h^{2-\delta})$
barrier	$O(h^{1/2})$	$o(h^{1/2-\delta})$	$O(h^{3/2})$	$o(h^{3/2-\delta})$
digital	$O(h^{1/2})$	$O(h^{1/2} \log h)$	$O(h^{3/2})$	$o(h^{3/2-\delta})$

Table:  $V_\ell$  convergence observed numerically (for GBM) and proved analytically (for more general SDEs)

Euler analysis due to G, Higham & Mao (2009) and Avikainen (2009).  
Milstein analysis due to G, Debrabant & Rößler (2012).

# Lévy processes

Infinite activity rate general Lévy processes

(Dereich 2010; Marxen 2010; Dereich & Heidenreich 2011)

- on level  $\ell$ , simulate jumps bigger than  $\delta_\ell$  ( $\delta_\ell \rightarrow 0$  as  $\ell \rightarrow \infty$ )
- either neglect smaller jumps or use a Gaussian approximation
- multilevel problem: discrepancy in treatment of jumps which are bigger than  $\delta_\ell$  but smaller than  $\delta_{\ell-1}$

# Lévy processes

## Exact simulation (Cheng Zhu, Filippo Zinzani, Yuan Xia)

- with some popular exponential-Lévy models (variance-gamma, NIG) possible to directly simulate Lévy increments over fine timesteps
- sum them pairwise to get corresponding increments for coarse path
- coarse and fine path simulations are both exact, so what's the point of multilevel simulation?
  - ▶ Asian options
  - ▶ lookback options
  - ▶ barrier options
  - ▶ other path-dependent options

# Heston stochastic volatility

Glasserman & Kim (2011) developed a series expansion for sampling from the integrated variance:

$$\left( \int_0^T V_s ds \mid V_0 = v_0, V_t = v_t \right) \stackrel{d}{=} \sum_{n=1}^{\infty} x_n + \sum_{n=1}^{\infty} y_n + \sum_{n=1}^{\infty} z_n$$

where  $x_n, y_n, z_n$  are independent random variables.

Multilevel possibility:

- truncate series at  $K_\ell$  ( $K_\ell \rightarrow \infty$  as  $\ell \rightarrow \infty$ )
- should help for European options as well as path-dependent options

# American options

Belomestny & Schoenmakers (2011) have developed a multilevel implementation of upper bound dual pricing

- based on nested simulation algorithm of Andersen and Broadie (2004)
- requires sub-sampling at each timestep to estimate a conditional expectation (the continuation value)
- multilevel treatment uses a different number of sub-samples  $M_\ell$  on each level ( $M_\ell \rightarrow \infty$  as  $\ell \rightarrow \infty$ )

- very natural straightforward application, with better savings than SDEs due to higher dimensionality
- big challenge is in numerical analysis – noteworthy contribution by Charrier, Scheichl & Teckentrup (2010)
- range of applications
  - ▶ Graubner & Ritter (2008) – parabolic
  - ▶ G, Reisinger (2009-11) – parabolic
  - ▶ Barth, Lang, Mishra, Schwab, Sukys, Zollinger (2010/11)  
– elliptic, parabolic, hyperbolic
  - ▶ Cliffe, G, Scheichl, Teckentrup (2010/11) – elliptic

# Engineering Uncertainty Quantification

- consider 3D elliptic PDE, with uncertain boundary data
- use grid spacing proportional to  $2^{-\ell}$  on level  $\ell$
- cost is  $O(2^{-3\ell})$ , if using an efficient multigrid solver
- 2nd order accuracy means that

$$\begin{aligned}\widehat{P}_\ell(\omega) - \widehat{P}(\omega) &\approx c(\omega) 2^{-2\ell} \\ \implies \widehat{P}_{\ell-1}(\omega) - \widehat{P}_\ell(\omega) &\approx 3c(\omega) 2^{-2\ell}\end{aligned}$$

- hence,  $\alpha=2$ ,  $\beta=4$ ,  $\gamma=3$
- cost is  $O(\varepsilon^{-2})$  to obtain  $\varepsilon$  RMS accuracy

# Elliptic SPDE

Elliptic PDE with random coefficient  $k(\mathbf{x}, \omega)$ :

$$-\nabla \cdot (k(\mathbf{x}, \omega) \nabla p(\mathbf{x}, \omega)) = 0, \quad \mathbf{x} \in D,$$

Model  $k$  as a lognormal random field, i.e.  $\log k$  is a Gaussian field with mean 0 and covariance function

$$R(\mathbf{x}, \mathbf{y}) = \sigma^2 \exp\left(-\|\mathbf{x} - \mathbf{y}\|_1 / \lambda\right)$$

Samples of  $\log k$  are provided by a Karhunen-Loève expansion:

$$\log k(\mathbf{x}, \omega) = \sum_{n=0}^{\infty} \sqrt{\theta_n} \xi_n(\omega) f_n(\mathbf{x}),$$

where  $\xi_n$  are iid unit Normal random variables.

# Elliptic SPDE

In multilevel treatment:

- different spatial grid resolution on each level
- truncate KL-expansion at different cutoffs  $K_\ell$

$$\log k_\ell(\mathbf{x}, \omega) = \sum_{n=0}^{K_\ell} \sqrt{\theta_n} \xi_n(\omega) f_n(\mathbf{x}),$$

- (more efficient ways of generating  $\log k_\ell$  use technique known as *circulant embedding*)

# Stochastic chemical reactions

In stochastic simulations, each reaction is a Poisson process with a rate which depends on the current concentrations.

In the “tau-leaping” method (Euler-Maruyama method) the reaction rates are frozen at the start of the timestep, so for each reaction sample from a Poisson process

$$P(\lambda \Delta t)$$

to determine the number of reactions in that timestep.

(As  $\lambda \Delta t \rightarrow \infty$ , the standard deviation becomes smaller relative to the mean, and it approaches the deterministic limit.)

# Stochastic chemical reactions

Anderson & Higham (2011) have developed a very efficient multilevel version of this algorithm – big savings because finest level usually has 1000's of timesteps.

Key challenge: how to couple coarse and fine path simulations?

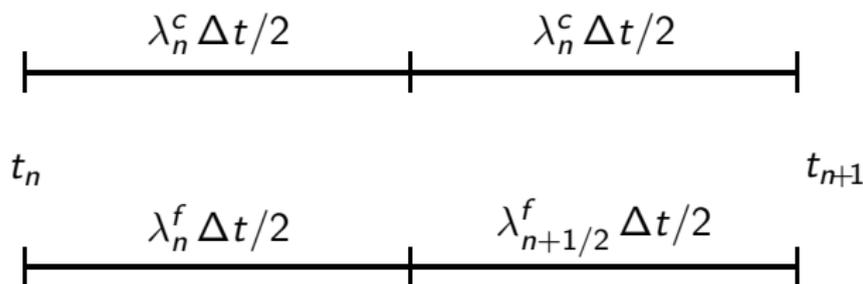
Crucial observation:  $P(t_1) + P(t_2) \stackrel{d}{=} P(t_1 + t_2)$

Only requirement:  $t_1, t_2 \geq 0$

# Stochastic chemical reactions

Solution:

- simulate the Poisson variable on the coarse timestep as the sum of two fine timestep Poisson variables
- couple the fine path and coarse path Poisson variables by using common variable based on smaller of two rates



If  $\lambda_n^f < \lambda_n^c$ , use  $P(\lambda_n^c \Delta t / 2) \sim P(\lambda_n^f \Delta t / 2) + P((\lambda_n^c - \lambda_n^f) \Delta t / 2)$

To further improve the multilevel complexity, can use randomised QMC in place of MC.

G & Waterhouse (2008-9) used rank-1 lattice rules for scalar SDE applications

- far fewer samples required on coarsest levels
- almost no difference on finest levels
- overall, big savings when using Milstein discretisation (so most work on coarsest levels)
- in best case (GBM with European option) complexity was approximately  $O(\varepsilon^{-1.5})$

Numerical algorithm:

- 1 start with  $L=0$
- 2 get an initial estimate for  $V_L$  using 32 random offsets and  $N_L = 1$
- 3 while  $\sum_{\ell=0}^L V_\ell > \varepsilon^2/2$ , try to maximise variance reduction per unit cost by doubling  $N_\ell$  on the level with largest  $V_\ell / (C_\ell N_\ell)$
- 4 if  $L < 2$  or the bias estimate is greater than  $\varepsilon/\sqrt{2}$ , set  $L := L+1$  and go back to step 2

# Conclusions

- multilevel idea is very simple; key is how to apply it in new situations
- lots of freedom to construct more efficient estimators:
  - ▶ change of measure
  - ▶ zero-mean control variate
  - ▶ antithetic treatment
  - ▶ sub-division
- being used for an increasingly wide range of applications; biggest computational savings when coarsest (helpful) approximation is much cheaper than finest
- currently, getting at least  $100\times$  savings for SPDEs and stochastic chemical reaction simulations

Webpage for my research/papers:

[people.maths.ox.ac.uk/gilesm/mlmc.html](http://people.maths.ox.ac.uk/gilesm/mlmc.html)

Webpage: [people.maths.ox.ac.uk/gilesm/mlmc\\_community.html](http://people.maths.ox.ac.uk/gilesm/mlmc_community.html)

Abo Academi (Avikainen) – numerical analysis  
Basel (Harbrecht) – elliptic SPDEs, sparse grid links  
Bath (Kyprianou, Scheichl, Shardlow) – elliptic SPDEs, MCMC, Lévy-driven SDEs  
Chalmers (Lang) – SPDEs  
Christian-Albrechts University (Gnewuch) – multilevel QMC  
Duisburg (Belomestny) – Bermudan and American options  
Edinburgh (Davie, Szpruch) – SDEs, numerical analysis  
ETH Zürich (Jenny, Jentzen, Schwab) – numerical analysis, SPDEs  
Frankfurt (Gerstner, Kloeden) – numerical analysis, sparse grid links  
Fraunhofer ITWM (Iliev) – SPDEs in engineering  
Hong Kong (Chen) – Brownian meanders, nested simulation in finance  
IIT Chicago (Hickernell) – SDEs, infinite-dimensional integration, complexity analysis  
Kaiserslautern (Heinrich, Korn, Ritter) – finance, SDEs, complexity analysis, parametric integration  
KAUST (Tempone) – adaptive time-stepping  
Kiel (Gnewuch) – randomized multilevel QMC  
Mannheim (Neuenkirch) – numerical analysis, fractional Brownian motion  
Marburg (Dereich) – Lévy-driven SDEs  
Munich (Hutzenthaler) – numerical analysis  
Oxford (Giles, Hambly, Reisinger) – SDEs, jump-diffusion, SPDEs, numerical analysis  
Passau (Müller-Gronbach) – infinite-dimensional integration, complexity analysis  
Purdue (Gittelsohn) – SDPEs  
Stanford (Glynn) – numerical analysis  
Strathclyde (Higham, Mao) – numerical analysis, exit times, stochastic chemical modelling  
Stuttgart (Barth) – SPDEs  
Texas A&M (Efendiev) – SPDEs in engineering  
UCLA (Caflisch) – Coulomb collisions in physics  
UNSW (Dick, Kuo, Sloan) – multilevel QMC  
WIAS (Schoenmakers) – Bermudan and American options  
Wisconsin (Anderson) – numerical analysis, stochastic chemical modelling

Three keynote talks (out of a total of eight):

- Steffen Dereich (WWU)
  - ▶ *Multilevel Monte Carlo for Lévy-driven SDEs*
  - ▶ looking at methods and analysis for MLMC for path-dependent functionals of Lévy-driven SDEs
- Peter Glynn (Stanford)
  - ▶ *Creating unbiased Monte Carlo schemes from biased ones: theory and applications*
  - ▶ modifies MLMC to randomise the selection of level for each sample
- Raúl Tempone (KAUST)
  - ▶ *Adaptive strategies for Multilevel Monte Carlo*
  - ▶ discusses adaptive timestepping for MLMC simulations

Also 4 MLMC sessions (3 invited, and 1 contributed on Thurs afternoon) with a total of 29 presentations covering a wide range of topics