# Multilevel Monte Carlo Path Simulation

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### **SDEs** in Finance

In computational finance, stochastic differential equations are used to model the behaviour of

- stocks
- interest rates
- exchange rates
- weather
- electricity/gas demand
- crude oil prices
- **\_**

The stochastic term accounts for the uncertainty of unpredictable day-to-day events.

### **SDEs** in Finance

These models are then used to calculate "fair" prices for a huge range of financial options:

- an option to sell a stock portfolio at a specific price in 2 years time
- an option to buy aviation fuel at a specific price in 6 months time
- an option to sell US dollars at a specific exchange rate in 3 years time

In most cases, the buyer of the financial option is trying to reduce their risk.

### **SDEs** in Finance

#### Examples:

 Geometric Brownian motion (Black-Scholes model for stock prices)

$$dS = r S dt + \sigma S dW$$

Cox-Ingersoll-Ross model (interest rates)

$$dr = \alpha(b - r) dt + \sigma \sqrt{r} dW$$

Heston stochastic volatility model (stock prices)

$$dS = r S dt + \sqrt{V} S dW_1$$
  

$$dV = \lambda (\sigma^2 - V) dt + \xi \sqrt{V} dW_2$$

with correlation  $\rho$  between  $dW_1$  and  $dW_2$ 

### **Generic Problem**

Stochastic differential equation with general drift and volatility terms: SDE with general drift and volatility terms:

$$dS(t) = a(S, t) dt + b(S, t) dW(t)$$

W(t) is a Wiener variable with the properties that for any q < r < s < t, W(t) - W(s) is Normally distributed with mean 0 and variance t-s, independent of W(r) - W(q).

In many finance applications, we want to compute the expected value of an option dependent on the terminal state

$$P = f(S(T))$$

with a uniform Lipschitz bound,

$$|f(U) - f(V)| \le c \|U - V\|, \quad \forall U, V.$$

# Standard MC Approach

Euler discretisation with timestep *h*:

$$\widehat{S}_{n+1} = \widehat{S}_n + a(\widehat{S}_n, t_n) h + b(\widehat{S}_n, t_n) \Delta W_n$$

Simplest estimator for expected payoff is an average of N independent path simulations:

$$\widehat{Y} = N^{-1} \sum_{i=1}^{N} f(\widehat{S}_{T/h}^{(i)})$$

# Standard MC Approach

#### Two kinds of errors:

statistical error, due to finite number of paths

$$V[\widehat{Y}] = N^{-1}V[f(\widehat{S}_{T/h})]$$

so r.m.s. error =  $O(N^{-1/2})$ .

- discretisation bias, due to finite number of timesteps
  - weak convergence O(h) error in expected payoff
  - strong convergence  $O(h^{1/2})$  error in individual path

# Standard MC Approach

Mean Square Error is  $O(N^{-1} + h^2)$ 

- first term comes from variance of estimator
- second term comes from bias due to weak convergence

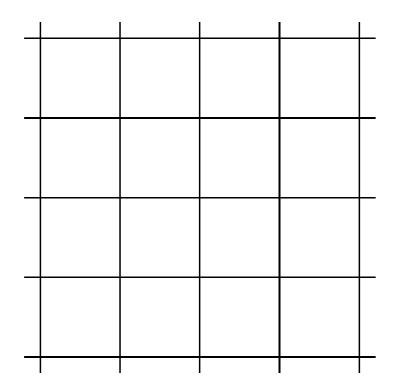
To make this  $O(\varepsilon^2)$  requires

$$N = O(\varepsilon^{-2}), \quad h = O(\varepsilon) \implies \cos t = O(N h^{-1}) = O(\varepsilon^{-3})$$

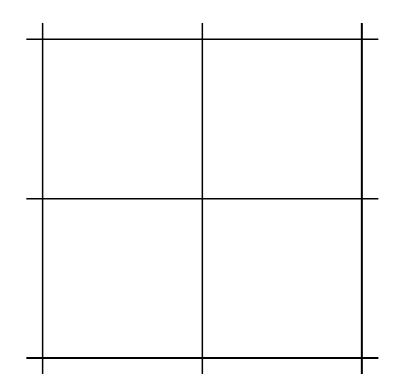
Aim is to improve this cost to  $O\left(\varepsilon^{-2}(\log \varepsilon)^2\right)$ 

# Multigrid

A powerful technique for solving PDE discretisations:



Fine grid more accurate more expensive



Coarse grid
less accurate
less expensive

# Multigrid

Multigrid combines calculations on a nested sequence of grids to get the accuracy of the finest grid at a much lower computational cost.

We will use a similar idea to achieve variance reduction in Monte Carlo path calculations, combining simulations with different numbers of timesteps – same accuracy as finest calculations, but at a much lower computational cost.

### **Other Research**

- In Dec. 2005, Ahmed Kebaier published an article in Annals of Applied Probability describing a two-level method which reduces the cost to  $O(\varepsilon^{-2.5})$ .
- Also in Dec. 2005, Adam Speight wrote a working paper describing a similar multilevel use of control variates, but without an analysis of its complexity.
- There are also close similarities to a multilevel technique developed by Stefan Heinrich for parametric integration (Journal of Complexity, 1998)

Consider multiple sets of simulations with different timesteps  $h_l = 2^{-l} T$ , l = 0, 1, ..., L, and payoff  $\widehat{P}_l$ 

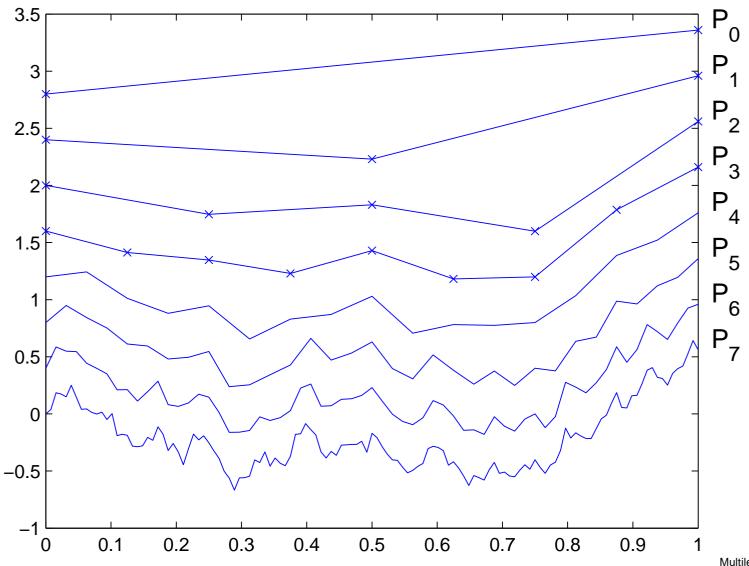
$$E[\widehat{P}_{L}] = E[\widehat{P}_{0}] + \sum_{l=1}^{L} E[\widehat{P}_{l} - \widehat{P}_{l-1}]$$

Expected value is same – aim is to reduce variance of estimator for a fixed computational cost.

Key point: approximate  $E[\widehat{P}_l - \widehat{P}_{l-1}]$  using  $N_l$  simulations with  $\widehat{P}_l$  and  $\widehat{P}_{l-1}$  obtained using <u>same</u> Brownian path.

$$\widehat{Y}_{l} = N_{l}^{-1} \sum_{i=1}^{N_{l}} \left( \widehat{P}_{l}^{(i)} - \widehat{P}_{l-1}^{(i)} \right)$$

#### Discrete Brownian path at different levels



- each level adds more detail to Brownian path
- $E[\widehat{P}_l \widehat{P}_{l-1}]$  reflects impact of that extra detail on the payoff
- different timescales handled by different levels
   similar to different wavelengths being handled by different grids in multigrid

Using independent paths for each level, the variance of the combined estimator is

$$V\left[\sum_{l=0}^{L} \widehat{Y}_{l}\right] = \sum_{l=0}^{L} N_{l}^{-1} V_{l}, \qquad V_{l} \equiv V[\widehat{P}_{l} - \widehat{P}_{l-1}],$$

and the computational cost is proportional to  $\sum_{l=0}^{L} N_l h_l^{-1}$ .

Hence, the variance is minimised for a fixed computational cost by choosing  $N_l$  to be proportional to  $\sqrt{V_l h_l}$ .

The constant of proportionality can be chosen so that the combined variance is  $O(\varepsilon^2)$ .

For the Euler discretisation and the Lipschitz payoff function

$$V[\widehat{P}_l - P] = O(h_l) \implies V[\widehat{P}_l - \widehat{P}_{l-1}] = O(h_l)$$

and the optimal  $N_l$  is asymptotically proportional to  $h_l$ .

To make the combined variance  $O(\varepsilon^2)$  requires

$$N_l = O(\varepsilon^{-2}L\,h_l).$$

To make the bias  $O(\varepsilon)$  requires

$$L = \log_2 \varepsilon^{-1} + O(1) \implies h_L = O(\varepsilon).$$

Hence, we obtain an  $O(\varepsilon^2)$  MSE for a computational cost which is  $O(\varepsilon^{-2}L^2) = O(\varepsilon^{-2}(\log \varepsilon)^2)$ .

**Theorem:** Let P be a functional of the solution of a stochastic o.d.e., and  $\widehat{P}_l$  the discrete approximation using a timestep  $h_l = M^{-l} T$ .

If there exist independent estimators  $\widehat{Y}_l$  based on  $N_l$  Monte Carlo samples, and positive constants  $\alpha \geq \frac{1}{2}, \beta, c_1, c_2, c_3$  such that

i) 
$$E[\widehat{P}_l - P] \le c_1 h_l^{\alpha}$$

ii) 
$$E[\widehat{Y}_l] = \begin{cases} E[\widehat{P}_0], & l = 0 \\ E[\widehat{P}_l - \widehat{P}_{l-1}], & l > 0 \end{cases}$$

iii) 
$$V[\widehat{Y}_l] \leq c_2 N_l^{-1} h_l^{\beta}$$

iv)  $C_l$ , the computational complexity of  $\widehat{Y}_l$ , is bounded by

$$C_l \le c_3 \, N_l \, h_l^{-1}$$

**then** there exists a positive constant  $c_4$  such that for any  $\varepsilon < e^{-1}$  there are values L and  $N_L$  for which the multi-level estimator

$$\widehat{Y} = \sum_{l=0}^{L} \widehat{Y}_l,$$

has Mean Square Error 
$$MSE \equiv E\left[\left(\widehat{Y} - E[P]\right)^2\right] < \varepsilon^2$$

with a computational complexity C with bound

$$C \le \begin{cases} c_4 \varepsilon^{-2}, & \beta > 1, \\ c_4 \varepsilon^{-2} (\log \varepsilon)^2, & \beta = 1, \\ c_4 \varepsilon^{-2 - (1 - \beta)/\alpha}, & 0 < \beta < 1. \end{cases}$$

#### Geometric Brownian motion:

$$dS = r S dt + \sigma S dW, \qquad 0 < t < 1,$$

$$S(0) = 1$$
,  $r = 0.05$ ,  $\sigma = 0.2$ 

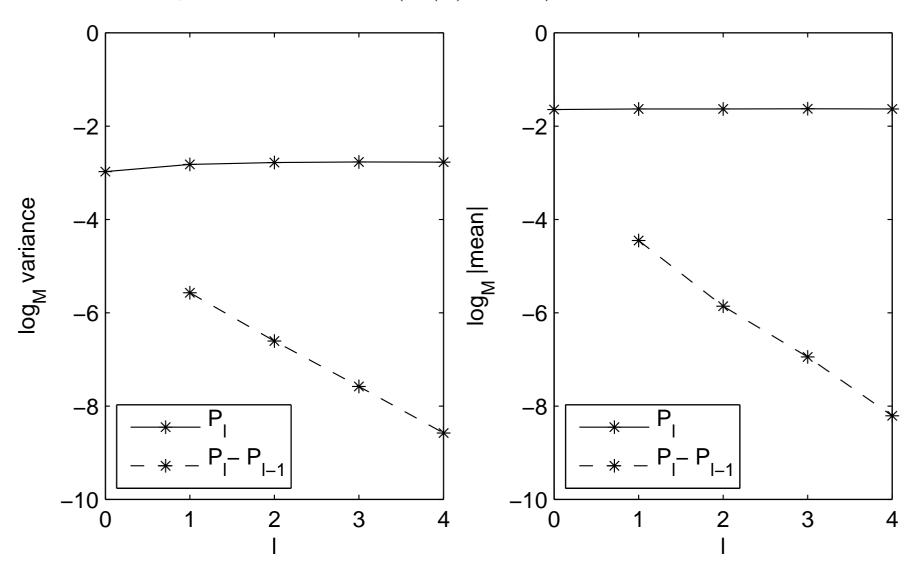
#### Heston model:

$$dS = r S dt + \sqrt{V} S dW_1, \qquad 0 < t < 1$$
  
$$dV = \lambda (\sigma^2 - V) dt + \xi \sqrt{V} dW_2,$$

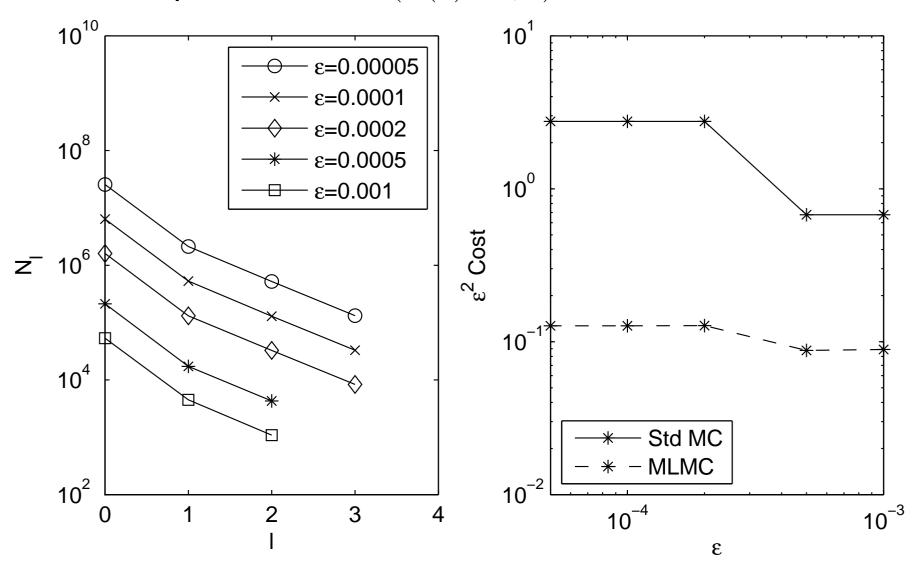
$$S(0) = 1$$
,  $V(0) = 0.04$ ,  $r = 0.05$ ,  $\sigma = 0.2$ ,  $\lambda = 5$ ,  $\xi = 0.25$ ,  $\rho = -0.5$ 

All calculations use M=4, more efficient than M=2.

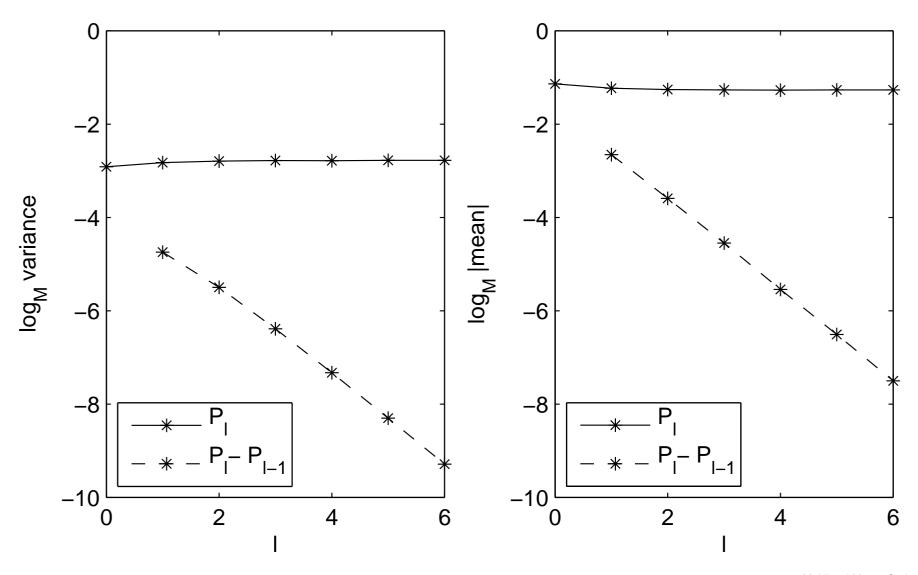
GBM: European call, max(S(1)-1,0)



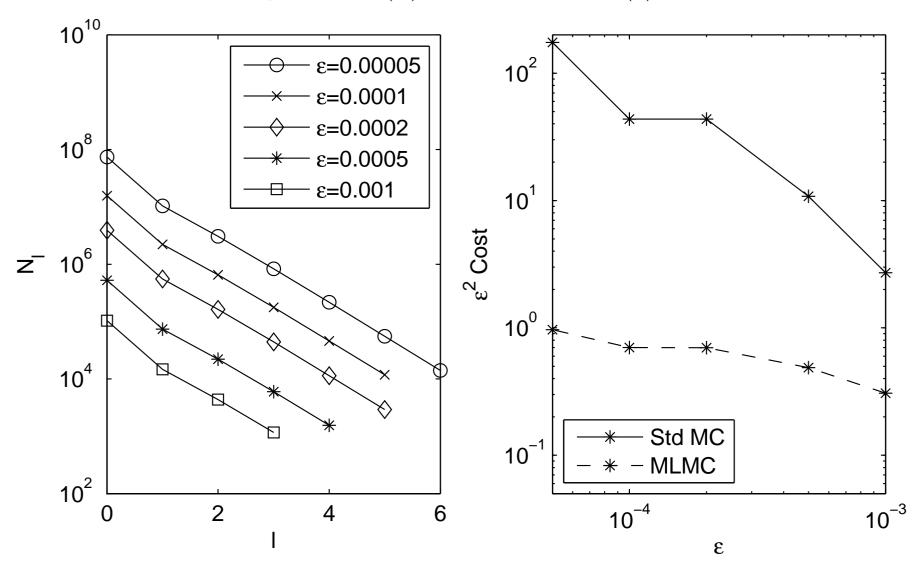
#### GBM: European call, max(S(1)-1,0)



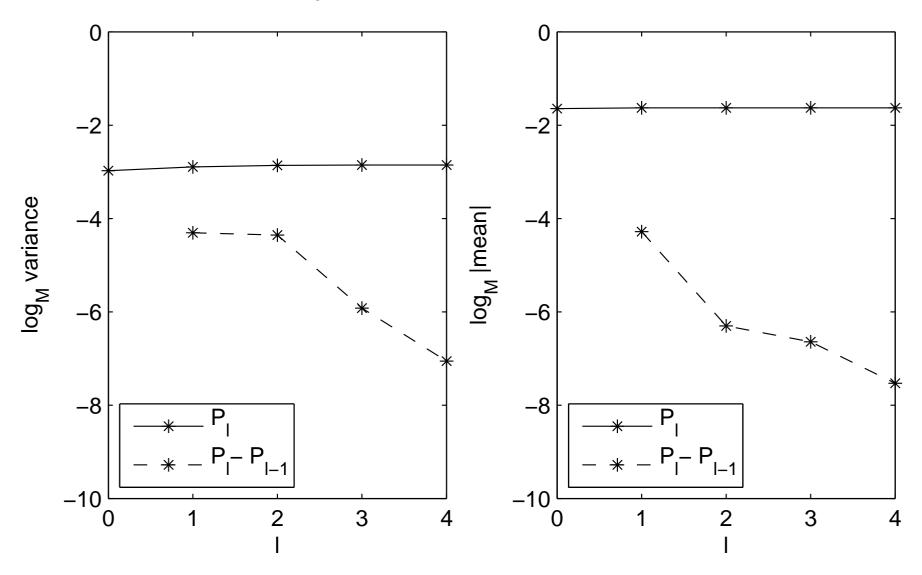
GBM: lookback option,  $S(1) - \min_{0 < t < 1} S(t)$ 



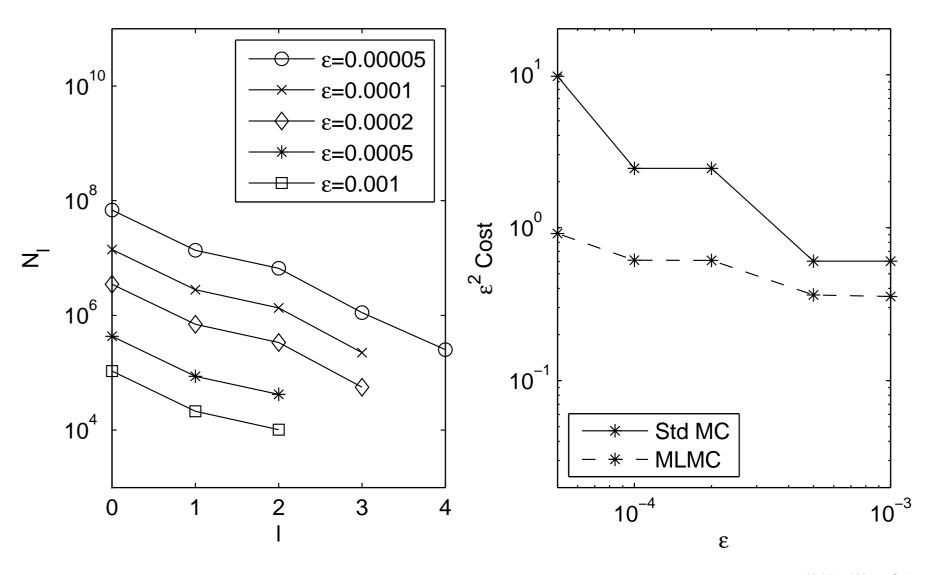
GBM: lookback option,  $S(1) - \min_{0 < t < 1} S(t)$ 



#### Heston model: European call



#### Heston model: European call



### **Conclusions**

#### Results so far:

- improved order of complexity
- easy to implement
- significant benefits for model problems

#### Current research:

- use of Milstein method (and antithetic variables in multi-dimensional case) to reduce complexity to  $O(\varepsilon^{-2})$
- adaptive sampling to treat discontinuous payoffs and pathwise derivatives for Greeks
- use of quasi-Monte Carlo methods, to reduce complexity towards  $O(\varepsilon^{-1})$

# **Working Paper**

M.B. Giles, "Multi-level Monte Carlo path simulation" Oxford University Computing Laboratory Numerical Analysis Report NA-06/03

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#### Generic SDE:

$$dS(t) = a(S, t) dt + b(S, t) dW(t), 0 < t < T,$$

with correlation matrix  $\Omega(S,t)$  between elements of  $\mathrm{d}W(t)$ .

Simplest Milstein scheme sets Lévy areas to zero to give

$$\widehat{S}_{i,n+1} = \widehat{S}_{i,n} + a_i h + b_{ij} \Delta W_{j,n} + \frac{1}{2} \frac{\partial b_{ij}}{\partial S_l} b_{lk} \left( \Delta W_{j,n} \Delta W_{k,n} - h \Omega_{jk} \right)$$

using implied summation convention.

#### In scalar case:

- O(h) strong convergence
- $O(\varepsilon^{-2})$  complexity for Lipschitz payoffs
- $O(\varepsilon^{-2})$  complexity for lookback, barrier and digital options using carefully constructed estimators

#### In multi-dimensional case:

- still only  $O(h^{1/2})$  strong convergence
- but  $\widehat{S}_n E[S \mid W_n] = O(h)$

If a coarse path with timestep 2h is constructed using

$$\Delta W_n^c = \sqrt{2h} \ Y_n$$

where the  $Y_n$  are N(0,1) random variables, and the fine path uses a Brownian Bridge construction with

$$\Delta W_n^f = \frac{1}{2}\sqrt{2h}(Y_n + Z_n), \quad \Delta W_{n+\frac{1}{2}}^f = \frac{1}{2}\sqrt{2h}(Y_n - Z_n).$$

where the  $Z_n$  are also N(0,1) random variables, then perturbation analysis shows that the  $O(h^{1/2})$  difference between the two paths comes from a sum of terms proportional to

$$Y_{j,n}Z_{k,n}-Y_{k,n}Z_{j,n}$$
.

Using the idea of antithetic variables, we use the estimator

$$\widehat{Y}_l = N_l^{-1} \sum_{i=1}^{N_l} \left( \frac{1}{2} \left( \widehat{P}_l^{(i)} + \widehat{P}_l^{(i)*} \right) - \widehat{P}_{l-1}^{(i)} \right),$$

where  $\widehat{P}_l^{(i)*}$  is based on the same coarse path  $Y_n$ , but with  $Z_n$  replaced by  $-Z_n$ , which leads to the cancellation of the leading order error proportional to  $Z_n$ .

- $V[\widehat{Y}_l] = O(h^2)$  for smooth payoffs,  $O(h^{3/2})$  for Lipschitz
- ${\color{red} \bullet}$  in both cases, gives  $O(\varepsilon^{-2})$  complexity for  $O(\varepsilon)$  accuracy

# **Adaptive sampling**

With digital options, the problem is that small path changes can lead to an  $\mathcal{O}(1)$  change in the payoff

For the Euler discretisation,  $O(h^{1/2})$  strong convergence  $\implies O(h^{1/2})$  paths have an O(1) value for  $\widehat{P}_l - \widehat{P}_{l-1}$ 

Hence,

$$V_l = O(h^{1/2}).$$

For improved results, need more samples of paths near payoff discontinuities.

# **Adaptive sampling**

Two ideas for adaptive sampling are both based on Brownian Bridge constructions, using coarse timestep realisations to decide which paths are important

- idea 1: start with relatively few paths, and sub-divide those which look interesting (splitting)
- idea 2: start with lots of paths, and prune those which are unimportant (Russian roulette)
- use path weights to ensure estimator remains unbiased
- initial results (combining 2 ideas to keep a fixed number of paths) look good for a digital option, and it should also handle barrier options

# **Quasi-Monte Carlo**

Quasi-Monte Carlo methods can offer greatly improved convergence with respect to the number of samples N:

- in the best case,  $O(N^{-1+\delta})$  error for arbitrary  $\delta > 0$ , instead of  $O(N^{-1/2})$
- depends on knowledge/identification of "important dimensions" in an application
  - Brownian Bridge
  - Principal Component Analysis
- confidence intervals can be obtained by using randomized QMC
- working with Sloan, Kuo and Waterhouse, will try both rank-1 lattice rules and Sobol sequences